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Striatal and Behavioral Responses to Reward Vary by Socioeconomic Status in Adolescents

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Citation: Alexandra L. Decker, Steven L. Meisler, Nicholas A. Hubbard, Clemens C.C. Bauer, Julia Leonard, Hannah Grotzinger, Melissa A. Giebler, Yesi Camacho Torres, Andrea Imhof, Rachel Romeo, John D. E. Gabrieli. Journal of Neuroscience 22 January 2024, e1633232023.

As Published: https://doi.org/10.1523/JNEUROSCI.1633-23.2023

Persistent URL: https://hdl.handle.net/1721.1/153394

Version: Author's final manuscript: final author's manuscript post peer review, without publisher's formatting or copy editing

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1	Striatal and Behavioral Responses to Reward Vary by Socioeconomic Status in
2	Adolescents
3	Abbreviated title: Reward Responses Vary By Adolescent Socioeconomic Status
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2:	5 Number of Pages: 39
20	Number of Figures: 6; Number of Tables: 0
2	7 Conflict of interest statement: the authors declare no competing interests
23 29	8 Competing interest statement: The authors declare no competing interests.
3	Key Words: Socioeconomic Status, Adolescence, Reward, Striatum, Decision-Making
3	1 Response Time
32	2
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Abstract

41 Disarities in socioeconomic status (SES) lead to unequal access to financial and social 42 support. These disparities are believed to influence reward sensitivity, which in turn, are 43 hypothesized to shape how individuals respond to and pursue rewarding experiences. 44 However, surprisingly little is known about how SES shapes reward sensitivity in 45 adolescence. Here we investigated how SES influenced adolescent responses to 46 reward, both in behavior and the striatum-a brain region that is highly sensitive to 47 reward. We examined responses to both immediate reward (tracked by phasic 48 dopamine) and average reward rate fluctuations (tracked by tonic dopamine) as these 49 distinct signals independently shape learning and motivation. Adolescents (n=114; 12-50 14 years; 58 female) performed a gambling task during functional magnetic resonance 51 imaging. We manipulated trial-by-trial reward and loss outcomes, leading to fluctuations 52 between periods of reward scarcity and abundance. We found that a higher reward rate 53 hastened behavioral responses, and increased guess switching, consistent with the 54 idea that reward abundance increases response vigor and exploration. Moreover, 55 immediate reward reinforced previously rewarding decisions (win-stay, lose-switch) and 56 slowed responses (post-reward pausing), particularly when rewards were scarce. 57 Notably, lower-SES adolescents slowed down less after rare rewards than higher-SES 58 adolescents. In the brain, striatal activations covaried with the average reward rate 59 across time, and showed greater activations during rewarding blocks. However, these 60 striatal effects were diminished in lower-SES adolescents. These findings show that the 61 striatum tracks reward rate fluctuations, which shape decisions and motivation.

62	Moreover, lower SES appears to attenuate reward-driven behavioral and brain
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Significance statement

Lower socioeconomic status (SES) is associated with reduced access to resources and opportunities. Such disparities may shape reward sensitivity, which in turn, could influence how individuals respond to and pursue rewarding experiences. Here, we show that lower-SES adolescents display reduced reward sensitivity in the brain and behavior. The striatum-a brain region that is highly sensitive to reward-showed greater activations during periods of high reward and tracked fluctuations between reward-rich and reward-scarce task phases. However, lower SES correlated with smaller reward-driven striatal responses, and reduced response slowing after rare rewards. These findings link lower SES to reduced reward responses, which could trigger a cycle of reduced reward pursuit, leading to fewer positive experiences, which could further diminish reward sensitivity.

Introduction

107 Adolescents from lower socioeconomic status (SES) backgrounds have less 108 access to enriching opportunities and resources than their higher-SES peers (Farah, 109 2017). These disparities may influence reward sensitivity, which in turn, could shape 110 how adolescents respond to or pursue rewarding experiences (Amir et al., 2018). Such 111 a cycle could explain how SES—by modulating reward responses and related 112 processes—is associated with many consequential developmental outcomes (Farah, 113 2017). Here, we examined how SES relates to reward-driven responses in behavior and 114 the brain in adolescents, focusing on the striatum because of its high sensitivity to 115 reward (Schultz, 1993). 116 Rewards powerfully influence motivation, learning, and decision-making. 117 Immediately rewarding outcomes, signaled by fast phasic striatal responses, are 118 thought to serve as a learning signal to maximize rewards (Day et al., 2007). Rewarding 119 outcomes strongly reinforce prior actions that led to rewards (Hamid et al., 2016) and 120 induce "post-reward pausing" in behavior (Schlinger et al., 2008). Individuals are also 121 sensitive to the overall amount of reward available in their environment. The average 122 environmental reward rate (tracked by tonic dopamine and estimated from past reward 123 history) influence moment-to-moment shifts in response time and exploration (Hamid et 124 al., 2016; Niv et al., 2007; Wang et al., 2021). A high environmental reward rate boosts 125 response speeding, in theory, by increasing the cost of time (slower responses forfeit

more rewards) (Beierholm et al., 2013; Niv et al., 2006, 2007; Otto & Daw, 2019; Wang

127 et al., 2013, 2021) and increases exploration, in theory, due to the high likelihood of

attaining rewards in the environment (Constantino & Daw, 2015; Niv et al., 2007;

129 Sukumar et al., 2021). Interestingly, these distinct reward signals also interact: reward

130 scarcity heightens sensitivity to immediate reward, amplifying both phasic dopamine

131 firing following rewards (Bayer & Glimcher, 2005; Hamid et al., 2016) and behavioral

132 pausing after rewarding outcomes (Schlinger et al., 2008).

133 How SES influences responses to these distinct reward signals in adolescents in 134 the brain and behavior remains unclear. Previous research suggests that lower SES 135 may increase sensitivity to immediate reward, as lower-SES individuals tend to choose 136 small immediate rewards over larger, delayed ones (Oshri et al., 2019). This is 137 hypothesized to adaptively enable individuals to quickly seize scarce reward 138 opportunities to meet basic needs (Frankenhuis et al., 2016; Frankenhuis & Nettle, 139 2020; Pepper & Nettle, 2017). Lower SES environments can also be less predictable (Evans, 2004), meaning past reward history may poorly predict future outcomes 140 141 (Behrens et al., 2007; Ross & Hill, 2002). Based on this research, lower-SES 142 adolescents may be highly responsive to *immediate* reward, but less responsive to past 143 reward history, which could lead to contextually sub-optimal behavior. 144 This hypothesis, however, contrasts with two studies that found that lower SES in 145 adolescents correlated with *reduced* responses to rewarding cues in the parietal (White 146 et al., 2022) and frontal (Palacios-Barrios et al., 2021) cortices. Notably, however, both 147 studies linked lower-SES to poorer behavioral learning of cue-reward associations 148 (Palacios-Barrios et al., 2021; statistical trend White et al., 2022), which may have

altered expectations of reward when viewing reward-predicting cues. The present studytherefore eliminated learning demands.

In the present study, we examined behavioral and striatal responses to reward and reward rate fluctuations in adolescents from diverse SES backgrounds. Adolescents performed a gambling task during functional magnetic resonance imaging in which they won or lost on each trial. Unbeknownst to participants, we manipulated trial outcomes, leading to alternating periods of reward scarcity and abundance. We examined how immediate reward and average reward rate fluctuations shaped vigor (response times or RTs) and choices differently by SES. We also examined SES-related differences in the influence of reward and average reward rate fluctuations on striatal responses. Our results support influential theories of decision that argue the striatum tracks average reward rate fluctuations, as well as theories that suggest that lower SES reduces behavioral and striatal reward sensitivity.

Methods

172 Participants

173 We recruited 127 adolescents from diverse SES backgrounds as part of a larger 174 project examining the relationship between SES, brain development, and cognition. 175 Eligible participants were in the 7th or 8th grade, proficient in English, had no MRI 176 contraindications, were not diagnosed with autism or a neurological disability, and were not born premature (<34 weeks). Thirteen children did not complete the MRI, resulting 177 178 in a sample of 114 adolescents (age range=12-14; mean (SD)=13.46 (0.68), n=56179 female). Five participants with excessive movement during scanning (average 180 framewise displacement (FD) of more than 0.6 mm) were retained only for behavioral 181 analyses, leaving 109 for the neuroimaging analysis (correlation between FD and SES 182 among the included participants: $\beta = 0.005$, SE = 0.01, t(107) = 0.35, p = .730, r = 0.03). 183 Of note, the findings remained unchanged with a more conservative limit of movement 184 (average FD of < 0.3mm). All children and their legal guardians provided assent and 185 consent. The study was approved by the MIT Committee on the Use of Human 186 Subjects. Participants received compensation for their time. 187 Before collecting data, we targeted a sample of at least 100 participants based 188 on studies reporting medium-to-large effects (i.e., Cohen's d of 0.5-0.8) on the 189 relationship between SES and cognitive performance (Finn et al., 2017; Leonard et al., 190 2019; Noble et al., 2007), brain structure (Decker, Duncan, et al., 2020; Romeo et al., 191 2018), and brain function (Finn et al., 2017). A sensitivity analysis revealed that our

192 sample size provided 80% power to detect medium effects (d of 0.53 or Pearson's *r* of
193 0.25) in two-tailed between subject analyses.

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195 Measure of socioeconomic status (SES)

196 Participants' caregivers reported their annual household income (range=\$2k-197 \$1.25m) and the number of years of schooling they had completed (*range*=7-20 years). 198 Our primary measure of SES incorporated both these variables. We averaged the z-199 scores of the maternal education, paternal education, and the z-score of the log-200 transformed income measures (Figure 1A depicts the SES distribution). The log 201 transformation on income accounts for the greater impact that gains have for lower-SES 202 individuals. Two participants were missing one of the three measures, so their SES 203 index was the average of the two others.

204

205 **Experimental Design**

206 Participants performed a variant of Delgado et al.'s card guessing task (Delgado 207 et al., 2000; Hubbard, Romeo, et al., 2020; Hubbard, Siless, et al., 2020); Figure 1B). 208 On each trial, adolescents guessed if an upcoming number, with a possible value from 209 1-9, would be larger or smaller than 5. They then received immediate feedback based 210 on the accuracy of their guess. Participants were told that accurate guesses would be 211 financially rewarded as wins, and inaccurate guesses would be financially punished as 212 losses, and that the sum of wins and losses would be calculated for an additional 213 payment. Unbeknownst to participants, trial-by-trial gains and losses were

predetermined, and fixed across trials, with numbers generated to match the
predetermined outcome for each trial. Outcomes were therefore unrelated to participant
guesses, which equalized uncertainty across participants, and ensured everyone had
the same experience of winning and losing.

218 Each trial began with a question mark, during which participants had 1.5 seconds 219 to register a guess (smaller than 5=index finger; larger than 5=middle finger; Figure 1B). 220 A number was then displayed for 500ms, followed by 500ms of feedback. Feedback 221 indicated whether participants had won or lost money, or neither won nor lost money. 222 Positive feedback, which followed correct guesses, consisted of a green arrow pointing 223 up and the text "+\$1"; negative feedback, which followed incorrect guesses, consisted of 224 a red arrow pointing down and text displaying "-\$0.5"; neutral feedback, which followed 225 the number 5, consisted of a light green double-sided arrow. If participants did not 226 register a guess, they received neutral feedback. This happened rarely (3.1 trials or 227 4.5% of trials on average per participant; relationship between missed responses and 228 SES: $\beta = 0.27$, SE = 0.29, t(114) = 0.95, p = .35). Participants viewed a fixation cross 229 for 1 second before a new trial began.

The task, in total, across both runs, consisted of 8 blocks of 8 trials each, with 4 blocks of mostly positive outcomes ('reward blocks') and 4 blocks of mostly negative outcomes ('loss blocks'). Each of the 2 runs contained 2 reward and 2 loss blocks and each block was approximately 28 seconds. This block design maximized the ability to detect striatal responses to reward, while also leading to alternating periods of monetary reward scarcity and abundance, allowing us to examine the influence of fluctuations in

average reward rate across time (Figure 1D). To keep participants unaware of the fixed
outcomes, there was no delay between blocks, and blocks contained a few trials of the
opposite type (Figure 1C depicts trial outcomes in a representative reward and loss
block). Reward blocks included 6 reward trials interleaved with two of either loss or
neutral trials. Loss blocks included 6 loss trials interleaved with two of either reward or
neutral trials. All participants received \$10 in bonus money after the task.



244 Figure 1. SES score distribution, gambling task schematic, and trial-by-trial

moving average of rewards and losses. (A) *Distribution of SES composite scores*:

SES was operationalized as the mean of the z-transformed maternal and paternal 246 247 education variables (years of schooling completed), and z-scored log of annual 248 household income. The distribution is displayed, with the y-axis representing sample 249 proportions. (B) Gambling task schematic: Participants guessed whether a forthcoming 250 number would be greater or less than 5. Next, the actual number was revealed, and 251 participants received positive (top panel, green arrow), negative (middle panel, red 252 arrow), or neutral feedback (if the number was 5; bottom panel, light green arrow) 253 regarding their guess. A 1-second fixation cross (not depicted) preceded the next trial. 254 (C) Calculating the moving average of reward: An exponentially weighted moving 255 average of gains and losses quantified recent history of reward. This measure was used 256 to examine whether distinct task phases associated with reward scarcity or abundance 257 influenced trial-by-trial shifts in behavior (choices, RTs), and interacted with behavioral 258 responses (choices, RTs) to immediate feedback. (D) Schematic of reward rate 259 fluctuations in a representative participant. Grey shading represents loss blocks and 260 non-grey areas represent reward blocks. Pink and green colors denote periods in which 261 the moving average of reward is above or below the mean.

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263 Image acquisition

- 264 Participants practiced the gambling task and completed a mock scanning session
- to acclimate to the MRI environment, which improves compliance (de Bie et al., 2010;
- Gao et al., 2023). They then completed 2 runs of the gambling task inside the scanner
- and watched a movie while we acquired a T1-weighted (T1w) anatomical scan. Images
- were acquired using a 3T Siemens Prisma Fit scanner with a 32-channel head coil.
- 269 Whole-brain functional BOLD images were acquired using an EPI sequence (TR=0.8s,
- 270 TE=37s, Flip Angle=52°, voxel size=2mm isotropic, multi-band factor=8). The two runs
- 271 were acquired with reversed phase encoding to support distortion correction. High
- 272 resolution T1w images were acquired with an MP-RAGE sequence (TR=2.4s, T
- 273 =2.18ms, Flip Angle=8°, voxel size=0.8mm isotropic).
- 274

275 Image preprocessing

276 Preprocessing of anatomical and functional data was performed using *fMRIPrep*277 version 22.1.1 (Esteban et al., 2019).

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279 Anatomical preprocessing

280 The anatomical T1w image was corrected for intensity non-uniformity with 281 N4BiasFieldCorrection (Tustison et al., 2010) distributed with ANTs 2.3.3 (Avants et al., 2008) and used as T1w-reference throughout the workflow. The T1w-reference was 282 283 then skull-stripped using ANTs workflow with OASIS30ANTs as target template. Brain 284 tissue segmentation of gray-matter, white-matter, and cerebrospinal fluid was performed 285 on the brain-extracted T1w using fast (FSL 6.0.5.1:57b01774, RRID:SCR_002823), 286 (Zhang et al., 2001). Brain surfaces were reconstructed using recon-all from FreeSurfer 287 version 7.2.0 (Dale et al., 1999)), and the brain mask estimated previously was refined 288 with a custom variation of the method to reconcile ANTs-derived and FreeSurfer-derived 289 segmentations of subcortical gray matter including striatal subregions (Fischl et al., 290 2002). Volume-based spatial normalization to one standard space was performed 291 through nonlinear registration, using brain-extracted versions of both the T1w reference 292 and T1w template. FSL's MNI ICBM 152 non-linear 6th Generation Asymmetric 293 Average Brain Stereotaxic Registration Model (Evans et al., 2012), RRID:SCR 002823; 294 TemplateFlow ID: MNI152NLin6Asym] was selected for spatial normalization. 295

296 Functional preprocessing

A skull-stripped reference volume was generated using a custom methodology of *fMRIPrep*. Head-motion parameters were estimated using mcflirt (FSL

299 6.0.5.1:57b01774, (Jenkinson et al., 2002). The estimated fieldmap was aligned with 300 rigid-body registration to the target EPI reference run. Field coefficients were mapped 301 onto the reference EPI using the rigid-body transform. BOLD runs were slice-time 302 corrected using 3dTshift from AFNI (Cox & Hyde, 1997) RRID:SCR_005927). The 303 BOLD reference images were co-registered to the T1w reference using bbregister 304 (FreeSurfer; (Greve & Fischl, 2009), with six degrees of freedom. Noise regressors 305 were estimated based on the preprocessed BOLD. Framewise displacement (FD) was 306 computed using two formulations following Power (Power et al., 2014), and Jenkinson 307 (Jenkinson et al., 2002). Physiological regressors were extracted from eroded 308 cerebrospinal fluid and white-matter volumes for use in subsequent, component-based 309 noise corrections (CompCor, Behzadi et al. 2007). The BOLD time-series were 310 resampled into standard space in a single interpolation step by composing all the 311 pertinent transformations (i.e., head-motion transform matrices, susceptibility distortion 312 correction, and co-registrations to anatomical and output spaces). Volumetric 313 resamplings were performed using ANTs, configured with Lanczos interpolation to 314 minimize the smoothing effects of other kernels (Lanczos 1964).

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316 Statistical analyses

317 Statistical analyses were conducted in R (version 4.2.2). Raw data, code, and 318 extended analyses and supplementary tables are available at the following link:

319 https://osf.io/pqtby/. Unless stated otherwise, linear mixed effects regressions or general 320 linear mixed effects regressions were employed for data that repeated within 321 participants (e.g., single-trial RTs). Mixed effects models included random intercepts for 322 each participant and random slopes for fixed effects that repeated within participants. In 323 case of non-converging models, we followed recommendations in (Brown, 2021). 324 iterating through the following until they converged: (1) using the "bobyga" optimizer, (2) 325 increasing the number of iterations, (3) forcing zero correlations among random effects, 326 and (4) dropping random effects based on model comparison. RTs that fell 3 absolute 327 deviations from an individual's median RT were excluded (n=2 on average per 328 participant). Measures were mean centered within or across participants or effect-coded 329 prior to model fitting. 330 331 Calculating trial-by-trial shifts in the moving average reward rate 332 We computed an exponentially weighted moving average (EWMA) of rewards 333 and losses across trials (\$1, \$0.5, or \$0; See *Figure 1C*-D for a schematic). Each trial 334 was assigned a value based on the recent reward and loss history. High values 335 indicated more gains than losses, whereas low values indicated more losses than gains. 336 We used an exponentially weighted (rather than simple) moving average to emphasize 337 recent time points, which have a larger impact on psychological state, while still 338 incorporating data points from farther in the past (Awheda & Schwartz, 2016). We used 339 the following update rule:

$$\mathsf{EWMA}_t = \alpha \times r_t + (1 - \alpha) \times \mathsf{EWMA}_{t-1}$$

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343	In this formula, the EWMAt represents the EWMA at the current trial, t, α is the
344	smoothing factor or learning rate parameter that determines the influence of the most
345	recent observation on the moving average, and r represents the reward on the current
346	trial, \underline{t} . To prevent disproportionate initial weighting and to ensure the average reward
347	rate stabilized, the first 3 trials were omitted from the EWMA measure. To balance
348	recent and historical data, α was based on an 8-period span, calculated as: $\alpha = 2/(N+1)$,
349	where N was set to 8 to match the number of trials in each fixed reward and loss block.
350	However, we found that using EWMAs derived from smoothing factors of 5- and 10-
351	period spans did not alter the pattern of results.

As an exploratory analysis, we also tested whether individual differences in optimal learning rates for the average reward rate variable differed by SES. To do so, we fit a model that estimated the learning rate as a free parameter for each participant using R's base *optim* function with the L-BFGS-B algorithm. The algorithm identified the learning rate per participant that minimized that residual sum of squares (RSS) in a model predicting subsequent RTs from the Exponentially Weighted Moving Average of Reward for each participant.

359 Characterizing behavioral responses to rewards

We examined how immediate feedback (win versus loss outcomes) and
 fluctuations in the average reward rate shaped RTs and guesses. We first fit a model

362 predicting RTs from the preceding trial's feedback (win, loss), the moving average of 363 reward, and their interaction. We then re-fit this model after adding SES as a covariate 364 and interaction term. We also examined the influence of immediate reward and average 365 reward rate fluctuations on choices-specifically, how likely an individual was to repeat 366 their prior quess or switch to a different quess (i.e., switched or stayed). Therefore, the 367 dependent variable was whether an individual had repeated their prior choice (switched 368 = 1; stayed = 0) and the independent variables were the preceding trial feedback (win, 369 loss), the moving average of reward, and their interaction. We re-fit this model after 370 adding SES as a covariate and interaction term. All models included trial number as a 371 covariate to control for the general effects on time on task. Since there were only 8 372 neutral trials per participant across the task, trials that followed neutral feedback were 373 excluded from analysis.

374

375 The relationship between SES and striatal volumes

Three linear mixed effects models were fit to examine the association between SES and ROI volumes, separately for the caudate, putamen, and nucleus accumbens. Each model predicted volume from SES, hemisphere, and their interaction, to determine whether the influence of SES was stronger for one particular hemisphere. Age, sex, and intracranial volume were also included as covariates. ROIs with volumes that fell > 3 absolute deviations from the sample median were excluded (all regions for 1 participant, the caudate and right nucleus accumbens for another).

384 Examining reward-driven striatal responses to reward and average reward rate 385 fluctuations across time

386 To ascertain if striatal activations during reward differed from loss blocks and to 387 examine their covariance with average reward rate fluctuations, we conducted 388 neuroimaging analyses with Nilearn. The scripts and data are publicly accessible 389 (https://osf.io/pqtby/). The approach involved two separate general linear models (GLMs) applied to participant data within MNI coordinate space. The first model had 390 391 distinct regressors for reward and loss conditions. The second model incorporated a 392 regressor for the exponentially weighted moving average of reward, resampled at the 393 fMRI's TR. Both models were convolved with SPM's hemodynamic response function 394 and controlled for head movement and noise components (3 translation and rotation 395 parameters, plus top 5 principal AcompCor components (defined in a combined white 396 matter and cerebrospinal fluid mask). This analysis yielded z-value effect size maps for 397 each subject. The maps were entered into a group-level analysis to identify striatal 398 voxels that were sensitive to the distinct reward versus loss blocks or to the average 399 reward rate. Sensitivity was defined by voxel significance within the anatomical striatal 400 mask from the Harvard-Oxford Atlas (FDR-corrected p < 0.05, minimum cluster size of 401 10). For each analysis, we calculated the mean z value per participant across 402 responsive voxels, separately for the caudate, putamen, and nucleus accumbens in 403 each hemisphere. Participants therefore had 6 z values (1 per ROI) for each analysis. 404 These values represented the average effect size for the differences in activations 405 between reward and loss blocks and the relationship with the average reward rate.

To assess the degree to which these effect sizes deviated from zero, we fit 2 intercept-only linear mixed-effects models, predicting mean z values per ROI, controlling for age and sex, with random intercepts per participant to account for repeated measures across hemispheres. We excluded outlier values that fell 3 absolute deviations from the sample's median (1 value for the left putamen and 1 for the left caudate). Including outliers did not change the pattern of results.

412 Examining how reward-driven striatal responses differ by SES

Finally, we tested how SES related to activation level differences between reward and loss blocks, as well as the degree to which striatal activations covaried with fluctuations in the average reward rate. To this end, we fit 2 linear mixed effects models. The dependent variables were z values reflecting, either, activation level differences for reward and loss blocks or the covariance between striatal activations and average reward rate fluctuations. Both models included SES, hemisphere, and their interaction as independent variables, and covariates for age and sex.

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Results

429 We first describe how behavioral responses, specifically response times (RT) 430 and choices, are influenced by immediately rewarding outcomes, and covary with 431 fluctuations in the average reward rate across time. We then describe how these 432 behavioral responses differ by SES. Turning to the neuroimaging data, we then explore 433 the association between SES and the volume of the putamen, caudate, and nucleus 434 accumbens. Furthermore, we examine differences in striatal activations during reward 435 versus loss blocks, and examine how these activations covary with temporal 436 fluctuations in the average reward rate. Finally, we focus on disparities in striatal 437 responses across SES. 438

439 Average reward rate fluctuations influence RTs and post-reward pausing

440 Adolescents responded more slowly after winning than losing (i.e., post-reward 441 pausing: $\beta = 0.02$, SE = 0.005, t(262) = 4.68, p < .001; Figure 2A). Furthermore, trial-by-442 trial RTs covaried with fluctuations in the average reward rate, such that a higher 443 average reward rate led to faster RTs ($\beta = -0.04$, SE = 0.02, t(98) = -2.32, p = .022). 444 Fluctuations in the average reward rate also interacted with immediate feedback to 445 shape RTs: periods of reward scarcity amplified post-reward pausing (Reward Rate x 446 Preceeding feedback: $\beta = -0.05$, SE = 0.02, t(104) = -3.53, p < .001, Figure 2B), 447 indicating responses to immediate reward were amplified by a history of low rewards. In 448 fact, post-reward pausing was only observed when rewards were scarce but not when 449 they were plentiful (effect of preceding feedback when the reward rate is centered at -

450 1SD below the mean: $\beta = 0.04$, SE = 0.007, t(100) = 6.23, p < .001; above the mean: β 451 = 0.003, SE = 0.008, t(104) = 0.43, p = .665). These findings show that adolescents 452 tracked fluctuations in the average reward rate, which shaped response times across 453 time and modulated sensitivity and responses to immediate reward.

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455





457 **choices.** (A) Adolescents responded more slowly after a win than a loss (p < 0.001). (B)

- 460 Adolescents were most likely to repeat a guess when their guess had been rewarded on
- the previous trial (p < 0.001). (D) These "post-reward stay" effects were amplified by a
- 462 low average reward rate (p < 0.001). In all figures, the mean and within-subject error
- 463 bars are plotted. The moving average reward rate was divided into low and high

⁴⁵⁸ A lower average reward rate amplified post-reward slowing (*interaction*: p < 0.001) 459 indicating heightened behavioral responses to reward when rewards were scarce. (C)

464 average reward rate using a median split for ease of visualization. Note that we model
465 the average reward rate continuously in all analyses.
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468 Average reward rate fluctuations influence guess switching

469 Immediate feedback reinforced decisions on subsequent trials: when adolescents 470 won, they were more likely to repeat their prior guess on the subsequent trial than if 471 they had lost ($\beta = -0.31$, SE = 0.04, z = -7.31, p < .001; Figure 2C). A lower average 472 reward rate also increased the likelihood of repeating a previously rewarded guess (i.e., 473 increased win-stay, lose-switch effects (Reward Rate x Preceeding Feedback: $\beta = 0.57$, 474 SE = 0.12, z = 4.73, p < .001; Figure 2D). Indeed, win-stay effects were most prominent 475 when the average reward rate was low, indicating a history of low rewards increased the 476 tendency to stick with a rare rewarding option (main effect of immediate feedback on choices when the average reward rate is centered at -1SD below the mean: $\beta = -0.51$, 477 478 SE = 0.05, z = -9.64, p < .001; above the mean: $\beta = -0.11$, SE = 0.06, z = -1.75, p =479 .080. In general, a history of high rewards (a higher average reward rate) also increased 480 the likelihood of switching guesses across trials ($\beta = 0.68$, SE = 0.15, z = 4.46, p < 481 .001), suggesting a greater tendency to make alternative exploratory decisions when 482 rewards were abundant. These findings suggest that a history of low reward increases 483 the tendency to stick with a previously rewarding option, and reduces the tendency to 484 explore alternatives for reward.

486 **Reward rate fluctuations influence post-reward pausing more in higher-SES**

487 adolescents

488 Immediate reward and average reward rate fluctuations influenced choices 489 similarly regardless of SES (SES x feedback: $\beta = 0.09$, SE = 0.05, z = 1.73, p = .084; 490 SES x average reward rate: $\beta = -0.25$, SE = 0.18, z = -1.40, p = .161; SES x feedback 491 type x moving average: $\beta = 0.18$, SE = 0.15, z = 1.21, p = .225; Figure 3A). Additionally, 492 these distinct temporal dimensions of reward influenced RTs similarly, regardless of 493 SES (SES x feedback: $\beta = -0.002$, SE = 0.005, t(109) = -0.45, p = .651; SES x average 494 reward rate: $\beta = 0.02$, SE = 0.02, t(100) = 1.04, p = .301; Figure 3A). 495 However, reward-rate fluctuations modulated post-reward pausing more in 496 higher- than lower-SES adolescents (SES x feedback type x moving average: β = 497 -0.04, SE = 0.02, t(105) = -2.54, p = .013; Figure 3B). That is, higher-SES adolescents 498 slowed more following rare rewards (main effect of SES when the reward rate is 499 centered at -1SD below the mean to reflect reward scarcity: $\beta = 0.02$, SE = 0.007, t(684) 500 = 2.20, p = .028; Figure 3B). When rewards were plentiful, higher-SES adolescents 501 slowed less following rewards than lower-SES adolescents (centered at +1SD above 502 the mean to reflect reward abundance: $\beta = -0.02$, SE = 0.008, t(4949) = -2.05, p = 503 .041; *Figure 3B*) though neither group showed significant evience of post-reward 504 pausing when rewards were plentiful (ps > 0.087). Interestingly, SES was unrelated to 505 individual differences in optimal learning rates ($\beta = 0.06$, SE = 0.04, t(114) = 1.32, p = 506 .189), suggesting that heightened post-reward pausing was not driven by a greater 507 tendency to update expectations in response to new information. These findings

- 508 suggest that adolescents from lower-SES backgrounds were less likely to adapt
- 509 responses to immediate reward based on average reward rate fluctuations. Analyses
- 510 reported in our extended analyses on the Open Science Framework
- 511 (<u>https://osf.io/9vhtw</u>) demonstrate these results are robust when using education and
- 512 income to separately characterize SES.
- 513
- A. The influence of reward on choices does not differ by SES



B. Reward rate fluctuations shape post-reinforcement pausing differently by SES



- 515 Figure 3. Reward rate fluctuations modulate post-reward pausing more in higher-
- 516 **SES adolescents**. (A) Reward rate fluctuations influenced choice switching following
- 517 immediate reward similarly by SES. (B) Reward rate fluctuations modulated post-reward

slowing following rewards when rewards were scarce among higher-SES adolescents.
In all figures, we depict the mean and within-subject error bars. The moving average
reward rate and socioeconomic status were divided into low and high bins using a
median split for ease of visualization. Note that we model these variables continuously
in the analyses reported in the paper.

pausing more among higher- than lower-SES adolescents. This led to greater RT

526 Lower SES correlates with smaller caudate volumes

- 527 Higher SES was associated with larger caudate volumes (β = 96.61, SE = 37.65,
- 528 t(103) = 2.57, p = .012; Figure 4). In contrast, there were no significant associations
- between SES and the volumes of the putamen (β = 32.78, SE = 49.74, t(103) = 0.66, p
- 530 = .511) or nucleus accumbens (β = 1.94, SE = 7.26, t(104) = 0.27, p = .790). Moreover,
- 531 there were no SES x hemisphere interactions in any ROI (all ps > 0.590), demonstrating
- 532 that SES-related differences in volumes did not differ by hemisphere.
- 533

518

Lower SES correlates with smaller caudate volumes





539 hemisphere as repeated measures within participants. Individual data points represent

- 540 participant level data, and the grey shading reflects the standard error of the mean. 541
- 542 The striatum tracks fluctuations in the average reward rate
- 543 Across adolescents, mean activations were larger during reward than loss blocks in the
- 544 caudate (β = 0.50, SE = 0.07, t(106) = 6.74, p < .001), putamen (β = 0.61, SE = 0.08,
- 545 t(106) = 7.83, p < .001), and nucleus accumbens ($\beta = 0.77$, SE = 0.08, t(107) = 9.87, p
- 546 <.001; *Figure 5A*). Furthermore, striatal activations covaried with the average reward
- 547 rate, such that a higher average reward rate led to greater activations in the caudate (β
- 548 = 0.77, SE = 0.08, t(105) = 9.60, p < .001), putamen ($\beta = 0.66$, SE = 0.07, t(105) = 8.94,
- 549 p < .001), and nucleus accumbens ($\beta = 1.32$, SE = 0.09, t(103) = 14.07, p < .001; Figure
- 550 5B. These findings show that the striatum not only responds more to reward than loss in
- 551 general, but tracks moment-by-moment shifts in the average reward rate across time.
- 552

A. Higher striatal activations in reward than loss blocks



B. The striatum tracks reward rate fluctuations across time



553

Figure 5. Striatal activations covary with reward and loss blocks and reward rate fluctuations across time. (A) Across the sample, activations in the striatum were greater during reward than loss blocks (ps < 0.001). (B) Moreover, the striatum tracked average reward rate fluctuations across time, even after FDR correction for multiple comparisons across every voxel in the striatum (ps < 0.001). In A and B, the color bar represents *z*-values and color intensity reflect the strength of the effect.

561

562 Lower SES correlates with reduced striatal responses to reward

563 Lower SES correlated with smaller activation level differences between reward

- and loss blocks in the caudate (β = 0.22, SE = 0.09, t(105) = 2.54, p = .013) and
- 565 putamen (β = 0.25, SE = 0.09, t(104) = 2.73, p = .007) and marginally in the nucleus
- 566 accumbens (marginal effect: β = 0.16, SE = 0.09, t(106) = 1.79, p = .077; Figure 6A).
- 567 None of these effects differed by hemisphere (SES x hemisphere: all ps > 0.29).
- 568 Furthermore, striatal activations covaried with average reward rate fluctuations more

strongly in higher-SES adolescents in the putamen ($\beta = 0.17$, SE = 0.09, t(104) = 2.02, p = .046; Figure 6B), but not the caudate ($\beta = 0.08$, SE = 0.10, t(104) = 0.88, p = .380) or nucleus accumbens ($\beta = 0.02$, SE = 0.11, t(101) = 0.18, p = .860). None of these effects differed by hemisphere (*SES x hemisphere*: all *p*s > 0.21). Of note, the relationship between SES and reward-driven activations also did not differ by striatal subregion (SES x subregion interaction: all ps > 0.10).





A. Lower SES correlates with smaller reward-driven activations

B. Putamen activations track reward rate fluctuations less in lower-SES adoelsecents





- 578 Lower SES correlated with less activation differences between reward and loss blocks
- 579 in the caudate and putamen and marginally in the nucleus accumbens. (B) The

relationship between striatal activations and average reward rate fluctuations was stronger for higher- than lower-SES adolescents in the putamen (p=0.046), but not in the caudate or nucleus accumbens (ps>0.37). For ease of visualization, individual data points reflect mean z values across the left and right hemisphere. Statistics reported in the text model the left and right hemisphere separately as repeated measures. Individual data points represent participant data, and the grey shading reflects the standard error of the mean.

- 587
- 588
- 589

Discussion

590 We asked how socioeconomic status in adolescence was related to reward-591 driven responses in the brain and behavior. Drawing on infuential models of decision-592 making (Constantino & Daw, 2015; Niv et al., 2006, 2007), we examined how choices, 593 response times, and striatal activations were shaped by immediate reward outcomes 594 and previous reward history (average reward rate fluctuations across time). We found 595 that, behaviorally, participants were more likely to repeat a guess if it had led to a win 596 (win-stay, lose-switch effects) and responded more slowly after receiving a reward 597 (post-reward pausing). Fluctuations in the average reward rate also shaped behavior: a 598 higher reward rate hastened RTs and increased guess switching. Moreover, a low 599 reward rate increased behavioral sensitivity to immediately rewarding outcomes, 600 augmenting win-stay, lose-switch effects and post-reward pausing. Notably, compared 601 to higher-SES adolescents, lower-SES adolescents exhibited reduced post-reward 602 pausing when rewards were scarce. We also observed that across participants, striatal 603 activations were larger during reward than loss blocks, and covaried with fluctuations in 604 the average reward rate across time. However, relative to higher-SES adolescents, 605 lower-SES adolescents displayed reduced activations during reward relative to loss

blocks in the caudate and putamen, and marginally in the nucleus accumbens. And,
putamen activations tracked average reward rate fluctuations less in lower-SES
adolescents. These findings show that the striatum tracks average reward rate
fluctuations, which shape choices and response times (Hamid et al., 2016; Niv et al.,
2006, 2007; Wang et al., 2013, 2021). They also link lower SES in adolescence to
reduced reward sensitivity, both in the brain and behavior.

We found that adolescents tracked fluctuations in the average reward rate across 612 613 time, which influenced decisions and response times. When rewards were abundant, 614 individuals were more likely to switch choices across trials. These findings align with 615 studies in human adults (Constantino & Daw, 2015; Niv et al., 2007; Sukumar et al., 616 2021) and support theories of decision making (Constantino & Daw, 2015; Sukumar et 617 al., 2021). These theories argue that when the average environmental reward rate is 618 *lower* that an option's perceived value, it is rational to 'stay' with a rewarding option due 619 to the limited prospects of finding rewards elsewhere. Conversely, when the 620 environmental reward rate is *higher* than the perceived value of an option, it makes 621 sense to switch to explore alternative sources of reward. It is possible, then, that 622 adolescents used the average reward rate as a threshold for whether to switch or stay 623 with a previous choice. Future research could examine how the tendency to track 624 average reward rate fluctuations develops-and whether adolescents-given their 625 heightened sensitivity to reward (Cohen et al., 2010; Davidow et al., 2016; Galvan, 626 2010; Galvan et al., 2006) might be even more attuned to fluctuations in the average 627 reward rate across time than adults.

628 A higher average reward rate also covaried with faster RTs. This finding is 629 consistent with research in human adults (Beierholm et al., 2013; Otto & Daw, 2019) 630 and supports theories arguing that fluctuations in the average reward rate shape the 631 cost time (Niv et al., 2006, 2007). That is, when rewards are abundant, action delays 632 are presumably more costly because one forfeits relatively more potential rewards. 633 incentivizing faster responses. Interestingly, other researchers have theorized that 634 rewards also govern the opportunity cost engaging effort and sustaining attention 635 (Esterman et al., 2016; Esterman & Rothlein, 2019; Kurzban et al., 2013; Lin et al., 636 2022; Massar et al., 2016; Otto & Daw, 2019) raising the possibility that average reward 637 rate fluctuations shape diverse aspects of cognition-such as fluctuations in attention 638 (Decker et al., 2023, 2022; Decker & Duncan, 2020). Our findings therefore not only 639 support theories linking reward rate fluctuations to motivation and decision-making and 640 extend these ideas to human adolescents, but raise questions about the influence of 641 reward rate fluctuations on other aspects of cognition.

642 Adolescents were also responsive to immediately rewarding outcomes, in line 643 with previous research (Hamid et al., 2016; Reynolds et al., 2001): they were most likely 644 to repeat a previous choice if it had led to a reward on the prior trial and responded 645 more slowly after a reward outcome, a phenomenon known as "post-reward pausing" 646 (Crossman, 1968; Felton & Lyon, 1966; McMillan, 1971; Schlinger et al., 2008; Wallace 647 & Mulder, 1973; Williams et al., 2011). Notably, these effects were amplified by a lower 648 average reward rate. Our finding adds to a growing body of research suggesting the 649 background average reward rate modulates sensitivity to immediate reward. Indeed, in

650 animals and humans, post-reward pausing is prolonged when rewards are scarce 651 (Schlinger et al., 2008). Furthermore, fewer recent rewards and lower tonic dopamine 652 amplify phasic dopamine firing (Hamid et al., 2016)-a finding that potentially provides a 653 neurobiological explanation for the increased rewards responsivity we observed here 654 when reward were scarce. Slower responses after unexpected reward could also reflect 655 surprise due to the infrequency of the event (Decker, Finn, et al., 2020) or heightened 656 response caution that facilitated more deliberate decision-making (Schlinger et al., 657 2008, p. 50). Altogether, this finding shows that average reward rate fluctuations 658 influenced responses to immediate outcomes, which shaped choices and RTs. When 659 adolescents tune into the average environmental reward rate, they may make more 660 adaptive decisions according to the overall rewards available in the environment. 661 We also observed that the extent of RT slowing after rare rewards varied by 662 SES. Adolescents from higher-SES backgrounds showed greater post-reward pausing 663 than lower-SES adolescents when rewards were scarce. This finding could reflect 664 greater attunement to reward rate fluctuations among higher-SES adolescents, which 665 would be expected to increase the saliency of receiving a rare reward when the reward 666 rate was low. However, exploratory analyses showed that SES did not correlate with 667 learning rates-the tendency to update the average reward rate in response to new 668 outcomes. Thus, greater post-reward pausing may instead reflect a greater responsivity 669 to rewards in reward-scarce contexts specifically, rather than a general tendency to 670 more readily update the average reward rate.

Interestingly, reward rate fluctuations covaried with striatal activations in the caudate, putamen and nucleus accumbens, such that a higher reward rate led to greater activations in these regions. These findings are consistent with animals studies showing that tonic dopamine fluctuations in the striatum track the average reward rate and in doing so shapes motivational vigor and decision-making (Hamid et al., 2016; Wang et al., 2013, 2021), and, as far as we know, is the first human fMRI study demonstrating this relationship.

678 Our results extend prior findings linking lower SES to diminished reward 679 sensitivity in *neocortical* regions like the anterior cingulate cortex (Palacios-Barrios et 680 al., 2021) and parietal cortex (White et al., 2022). Indeed, we observed that the extent of 681 reward-driven activations in the striatum differed by SES. Higher-SES adolescents 682 showed greater reward-driven activations than lower-SES adolescents in the putamen. 683 caudate, and marginally in the nucleus accumbens. Moreover, putamen activations 684 tracked average reward rate fluctuations less in lower-SES adolescents. Notably, prior 685 studies employed incremental learning tasks in which adolescents learned the value of 686 cues in predicting outcomes over time. Our focus on a reward task that did not involve 687 learnable cue-outcome contingencies broadens the literature by showing that reduced 688 reward sensitivity is even observed when eliminating learning demands.

689 Our findings support proposals that lower SES environments reduce reward 690 sensitivity (Seligman, 1972). Past literature suggests that chronic stress diminishes the 691 belief that actions have consequences rendering individuals less motivated to pursue 692 rewarding outcomes (Seligman, 1972). It is therefore possible that chronic stress and

reduced perceived control, which are more common among lower-SES individuals
(Farah, 2018; Hackman et al., 2010; Hackman & Farah, 2009; McLaughlin et al., 2014)
mediated the effects we observed here. Targeted research that employs direct measures
of stress could directly test this mechanism.

697 The present findings offer insights into why cognitive performance (Noble et al., 698 2007) and emotional well-being (Reiss, 2013) are often reduced in lower SES 699 adolescents. Reward sensitivity plays a vital role in many aspects of cognition, 700 influencing everything from the ability to learn important associations (Davidow et al., 701 2016) to the ability to remain attentive to important events (Esterman & Rothlein, 2019; 702 Shenhav et al., 2013). Rewards boost motivation (Frömer et al., 2021; Schultz, 1993; 703 Westbrook et al., 2021; Westbrook & Braver, 2016), and support success in short and 704 long-term endeavors, such as academic and workplace pursuits. Disparities in reward 705 sensitivity, therefore, may contribute to disparities in learning, attentional performance, 706 and motivation. Given the intimate link between reward sensitivity and emotional well-707 being, reduced reward sensitivity may contribute to the higher rates of depression 708 (Auerbach et al., 2022; Reiss, 2013) and lower life satisfaction observed in lower-SES 709 groups (Kahneman & Deaton, 2010). On a broader level, these insights stress the 710 importance of socioeconomic policies (Farah, 2018) aimed at reducing the burdens of 711 poverty to foster cognitive and emotional well-being in society.

712

713 Data availability statement: Code and data can be found at the following link:
 714 <u>https://osf.io/pqtby/</u>

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721	Visualization: A.D.; Writing – Original Draft: A.D.; Writing – Review and Editing: all
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724	Acknowledgements: We thank Hause Lin and Kenneth Harris for providing feedback
725	on the neuroimaging analyses. This research was supported by the William and Flora
726	Hewlett Foundation (#4429 [J. D. E. G.]) and a Natural Sciences and Engineering
727	Research Council of Canada Postdoctoral Fellowship [A.L.D.]
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