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Striatal and Behavioral Responses to Reward Vary by Socioeconomic Status in Adolescents

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1 **Striatal and Behavioral Responses to Reward Vary by Socioeconomic Status in**
2 **Adolescents**

3 Abbreviated title: Reward Responses Vary By Adolescent Socioeconomic Status

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31 Response Time

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Abstract

40
41 Disparities in socioeconomic status (SES) lead to unequal access to financial and social
42 support. These disparities are believed to influence reward sensitivity, which in turn, are
43 hypothesized to shape how individuals respond to and pursue rewarding experiences.
44 However, surprisingly little is known about how SES shapes reward sensitivity in
45 adolescence. Here we investigated how SES influenced adolescent responses to
46 reward, both in behavior and the striatum—a brain region that is highly sensitive to
47 reward. We examined responses to both immediate reward (tracked by phasic
48 dopamine) and average reward rate fluctuations (tracked by tonic dopamine) as these
49 distinct signals independently shape learning and motivation. Adolescents (n=114; 12-
50 14 years; 58 female) performed a gambling task during functional magnetic resonance
51 imaging. We manipulated trial-by-trial reward and loss outcomes, leading to fluctuations
52 between periods of reward scarcity and abundance. We found that a higher reward rate
53 hastened behavioral responses, and increased guess switching, consistent with the
54 idea that reward abundance increases response vigor and exploration. Moreover,
55 immediate reward reinforced previously rewarding decisions (win-stay, lose-switch) and
56 slowed responses (post-reward pausing), particularly when rewards were scarce.
57 Notably, lower-SES adolescents slowed down less after rare rewards than higher-SES
58 adolescents. In the brain, striatal activations covaried with the average reward rate
59 across time, and showed greater activations during rewarding blocks. However, these
60 striatal effects were diminished in lower-SES adolescents. These findings show that the
61 striatum tracks reward rate fluctuations, which shape decisions and motivation.

62 Moreover, lower SES appears to attenuate reward-driven behavioral and brain
63 responses.

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Significance statement

Lower socioeconomic status (SES) is associated with reduced access to resources and opportunities. Such disparities may shape reward sensitivity, which in turn, could influence how individuals respond to and pursue rewarding experiences. Here, we show that lower-SES adolescents display reduced reward sensitivity in the brain and behavior. The striatum—a brain region that is highly sensitive to reward—showed greater activations during periods of high reward and tracked fluctuations between reward-rich and reward-scarce task phases. However, lower SES correlated with smaller reward-driven striatal responses, and reduced response slowing after rare rewards. These findings link lower SES to reduced reward responses, which could trigger a cycle of reduced reward pursuit, leading to fewer positive experiences, which could further diminish reward sensitivity.

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Introduction

107 Adolescents from lower socioeconomic status (SES) backgrounds have less
108 access to enriching opportunities and resources than their higher-SES peers (Farah,
109 2017). These disparities may influence reward sensitivity, which in turn, could shape
110 how adolescents respond to or pursue rewarding experiences (Amir et al., 2018). Such
111 a cycle could explain how SES—by modulating reward responses and related
112 processes—is associated with many consequential developmental outcomes (Farah,
113 2017). Here, we examined how SES relates to reward-driven responses in behavior and
114 the brain in adolescents, focusing on the striatum because of its high sensitivity to
115 reward (Schultz, 1993).

116 Rewards powerfully influence motivation, learning, and decision-making.
117 Immediately rewarding outcomes, signaled by fast phasic striatal responses, are
118 thought to serve as a learning signal to maximize rewards (Day et al., 2007). Rewarding
119 outcomes strongly reinforce prior actions that led to rewards (Hamid et al., 2016) and
120 induce “post-reward pausing” in behavior (Schlinger et al., 2008). Individuals are also
121 sensitive to the overall amount of reward available in their environment. The average
122 environmental reward rate (tracked by tonic dopamine and estimated from past reward
123 history) influence moment-to-moment shifts in response time and exploration (Hamid et
124 al., 2016; Niv et al., 2007; Wang et al., 2021). A high environmental reward rate boosts
125 response speeding, in theory, by increasing the cost of time (slower responses forfeit
126 more rewards) (Beierholm et al., 2013; Niv et al., 2006, 2007; Otto & Daw, 2019; Wang
127 et al., 2013, 2021) and increases exploration, in theory, due to the high likelihood of

128 attaining rewards in the environment (Constantino & Daw, 2015; Niv et al., 2007;
129 Sukumar et al., 2021). Interestingly, these distinct reward signals also interact: reward
130 scarcity heightens sensitivity to immediate reward, amplifying both phasic dopamine
131 firing following rewards (Bayer & Glimcher, 2005; Hamid et al., 2016) and behavioral
132 pausing after rewarding outcomes (Schlinger et al., 2008).

133 How SES influences responses to these distinct reward signals in adolescents in
134 the brain and behavior remains unclear. Previous research suggests that lower SES
135 may increase sensitivity *to immediate* reward, as lower-SES individuals tend to choose
136 small immediate rewards over larger, delayed ones (Oshri et al., 2019). This is
137 hypothesized to adaptively enable individuals to quickly seize scarce reward
138 opportunities to meet basic needs (Frankenhuis et al., 2016; Frankenhuis & Nettle,
139 2020; Pepper & Nettle, 2017). Lower SES environments can also be less predictable
140 (Evans, 2004), meaning past reward history may poorly predict future outcomes
141 (Behrens et al., 2007; Ross & Hill, 2002). Based on this research, lower-SES
142 adolescents may be highly responsive to *immediate* reward, but less responsive to past
143 reward history, which could lead to contextually sub-optimal behavior.

144 This hypothesis, however, contrasts with two studies that found that lower SES in
145 adolescents correlated with *reduced* responses to rewarding cues in the parietal (White
146 et al., 2022) and frontal (Palacios-Barrios et al., 2021) cortices. Notably, however, both
147 studies linked lower-SES to poorer behavioral learning of cue-reward associations
148 (Palacios-Barrios et al., 2021; statistical trend White et al., 2022), which may have

149 altered expectations of reward when viewing reward-predicting cues. The present study
150 therefore eliminated learning demands.

151 In the present study, we examined behavioral and striatal responses to reward and
152 reward rate fluctuations in adolescents from diverse SES backgrounds. Adolescents
153 performed a gambling task during functional magnetic resonance imaging in which they
154 won or lost on each trial. Unbeknownst to participants, we manipulated trial outcomes,
155 leading to alternating periods of reward scarcity and abundance. We examined how
156 immediate reward and average reward rate fluctuations shaped vigor (response times
157 or RTs) and choices differently by SES. We also examined SES-related differences in
158 the influence of reward and average reward rate fluctuations on striatal responses. Our
159 results support influential theories of decision that argue the striatum tracks average
160 reward rate fluctuations, as well as theories that suggest that lower SES reduces
161 behavioral and striatal reward sensitivity.

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171 Methods

172 Participants

173 We recruited 127 adolescents from diverse SES backgrounds as part of a larger
174 project examining the relationship between SES, brain development, and cognition.
175 Eligible participants were in the 7th or 8th grade, proficient in English, had no MRI
176 contraindications, were not diagnosed with autism or a neurological disability, and were
177 not born premature (<34 weeks). Thirteen children did not complete the MRI, resulting
178 in a sample of 114 adolescents (age range=12-14; mean (SD)=13.46 (0.68), $n=56$
179 female). Five participants with excessive movement during scanning (average
180 framewise displacement (FD) of more than 0.6 mm) were retained only for behavioral
181 analyses, leaving 109 for the neuroimaging analysis (*correlation between FD and SES*
182 among the included participants: $\beta = 0.005$, $SE = 0.01$, $t(107) = 0.35$, $p = .730$, $r = 0.03$).
183 Of note, the findings remained unchanged with a more conservative limit of movement
184 (average FD of < 0.3mm). All children and their legal guardians provided assent and
185 consent. The study was approved by the MIT Committee on the Use of Human
186 Subjects. Participants received compensation for their time.

187 Before collecting data, we targeted a sample of at least 100 participants based
188 on studies reporting medium-to-large effects (i.e., Cohen's d of 0.5-0.8) on the
189 relationship between SES and cognitive performance (Finn et al., 2017; Leonard et al.,
190 2019; Noble et al., 2007), brain structure (Decker, Duncan, et al., 2020; Romeo et al.,
191 2018), and brain function (Finn et al., 2017). A sensitivity analysis revealed that our

192 sample size provided 80% power to detect medium effects (d of 0.53 or Pearson's *r* of
193 0.25) in two-tailed between subject analyses.

194

195 **Measure of socioeconomic status (SES)**

196 Participants' caregivers reported their annual household income (*range*=\$2k-
197 \$1.25m) and the number of years of schooling they had completed (*range*=7-20 years).

198 Our primary measure of SES incorporated both these variables. We averaged the z-
199 scores of the maternal education, paternal education, and the z-score of the log-
200 transformed income measures (*Figure 1A depicts the SES distribution*). The log
201 transformation on income accounts for the greater impact that gains have for lower-SES
202 individuals. Two participants were missing one of the three measures, so their SES
203 index was the average of the two others.

204

205 **Experimental Design**

206 Participants performed a variant of Delgado et al.'s card guessing task (Delgado
207 et al., 2000; Hubbard, Romeo, et al., 2020; Hubbard, Siless, et al., 2020); *Figure 1B*).

208 On each trial, adolescents guessed if an upcoming number, with a possible value from
209 1-9, would be larger or smaller than 5. They then received immediate feedback based
210 on the accuracy of their guess. Participants were told that accurate guesses would be
211 financially rewarded as wins, and inaccurate guesses would be financially punished as
212 losses, and that the sum of wins and losses would be calculated for an additional
213 payment. Unbeknownst to participants, trial-by-trial gains and losses were

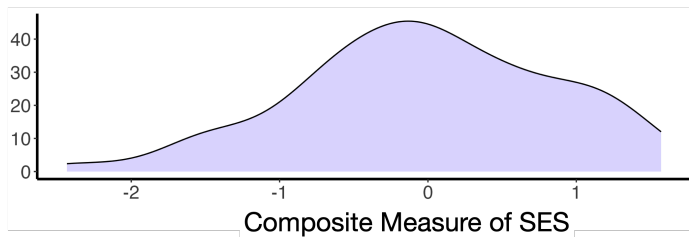
214 predetermined, and fixed across trials, with numbers generated to match the
215 predetermined outcome for each trial. Outcomes were therefore unrelated to participant
216 guesses, which equalized uncertainty across participants, and ensured everyone had
217 the same experience of winning and losing.

218 Each trial began with a question mark, during which participants had 1.5 seconds
219 to register a guess (smaller than 5=index finger; larger than 5=middle finger; Figure 1B).
220 A number was then displayed for 500ms, followed by 500ms of feedback. Feedback
221 indicated whether participants had won or lost money, or neither won nor lost money.
222 Positive feedback, which followed correct guesses, consisted of a green arrow pointing
223 up and the text “+\$1”; negative feedback, which followed incorrect guesses, consisted of
224 a red arrow pointing down and text displaying “-\$0.5”; neutral feedback, which followed
225 the number 5, consisted of a light green double-sided arrow. If participants did not
226 register a guess, they received neutral feedback. This happened rarely (3.1 trials or
227 4.5% of trials on average per participant; relationship between missed responses and
228 SES: $\beta = 0.27$, $SE = 0.29$, $t(114) = 0.95$, $p = .35$). Participants viewed a fixation cross
229 for 1 second before a new trial began.

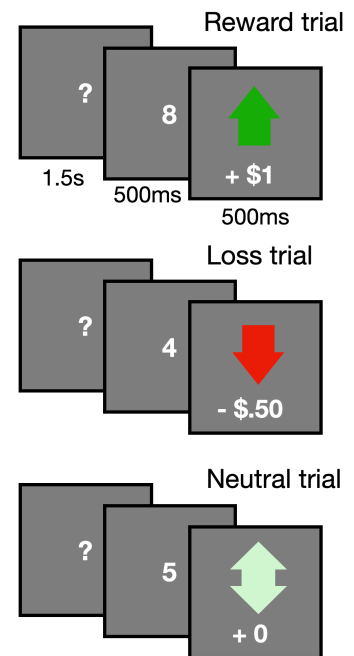
230 The task, in total, across both runs, consisted of 8 blocks of 8 trials each, with 4
231 blocks of mostly positive outcomes ('reward blocks') and 4 blocks of mostly negative
232 outcomes ('loss blocks'). Each of the 2 runs contained 2 reward and 2 loss blocks and
233 each block was approximately 28 seconds. This block design maximized the ability to
234 detect striatal responses to reward, while also leading to alternating periods of monetary
235 reward scarcity and abundance, allowing us to examine the influence of fluctuations in

236 average reward rate across time (Figure 1D). To keep participants unaware of the fixed
 237 outcomes, there was no delay between blocks, and blocks contained a few trials of the
 238 opposite type (Figure 1C depicts trial outcomes in a representative reward and loss
 239 block). Reward blocks included 6 reward trials interleaved with two of either loss or
 240 neutral trials. Loss blocks included 6 loss trials interleaved with two of either reward or
 241 neutral trials. All participants received \$10 in bonus money after the task.
 242

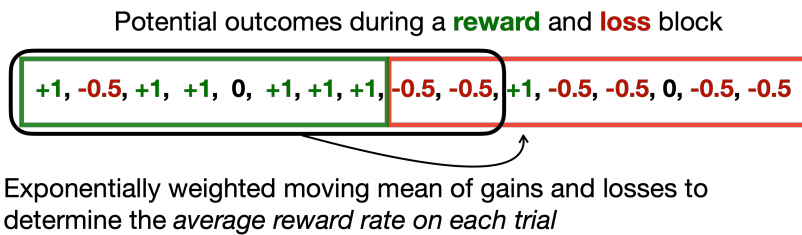
A. Distribution of Socioeconomic Status



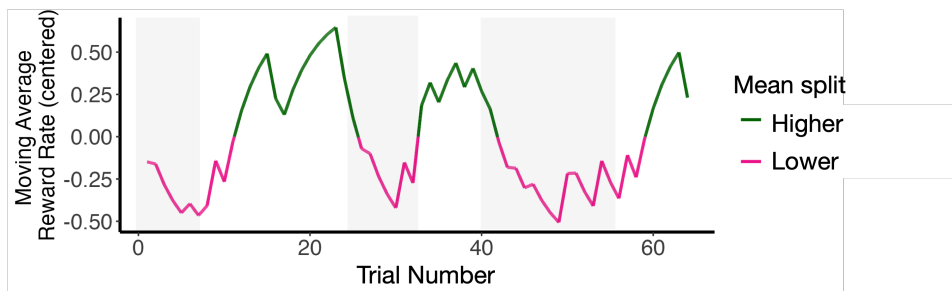
B. Gambling task schematic



C. Calculating the moving average of reward



D. Reward rate fluctuations in a representative participant



243

244 **Figure 1. SES score distribution, gambling task schematic, and trial-by-trial**
 245 **moving average of rewards and losses.** (A) Distribution of SES composite scores:

246 SES was operationalized as the mean of the z-transformed maternal and paternal
247 education variables (years of schooling completed), and z-scored log of annual
248 household income. The distribution is displayed, with the y-axis representing sample
249 proportions. (B) *Gambling task schematic*: Participants guessed whether a forthcoming
250 number would be greater or less than 5. Next, the actual number was revealed, and
251 participants received positive (top panel, green arrow), negative (middle panel, red
252 arrow), or neutral feedback (if the number was 5; bottom panel, light green arrow)
253 regarding their guess. A 1-second fixation cross (not depicted) preceded the next trial.
254 (C) *Calculating the moving average of reward*: An exponentially weighted moving
255 average of gains and losses quantified recent history of reward. This measure was used
256 to examine whether distinct task phases associated with reward scarcity or abundance
257 influenced trial-by-trial shifts in behavior (choices, RTs), and interacted with behavioral
258 responses (choices, RTs) to immediate feedback. (D) Schematic of reward rate
259 fluctuations in a representative participant. Grey shading represents loss blocks and
260 non-grey areas represent reward blocks. Pink and green colors denote periods in which
261 the moving average of reward is above or below the mean.
262

263 **Image acquisition**

264 Participants practiced the gambling task and completed a mock scanning session
265 to acclimate to the MRI environment, which improves compliance (de Bie et al., 2010;
266 Gao et al., 2023). They then completed 2 runs of the gambling task inside the scanner
267 and watched a movie while we acquired a T1-weighted (T1w) anatomical scan. Images
268 were acquired using a 3T Siemens Prisma Fit scanner with a 32-channel head coil.
269 Whole-brain functional BOLD images were acquired using an EPI sequence (TR=0.8s,
270 TE=37s, Flip Angle=52°, voxel size=2mm isotropic, multi-band factor=8). The two runs
271 were acquired with reversed phase encoding to support distortion correction. High
272 resolution T1w images were acquired with an MP-RAGE sequence (TR=2.4s, T
273 =2.18ms, Flip Angle=8°, voxel size=0.8mm isotropic).

274

275 **Image preprocessing**

276 Preprocessing of anatomical and functional data was performed using *fMRIPrep*
277 version 22.1.1 (Esteban et al., 2019).

278

279 *Anatomical preprocessing*

280 The anatomical T1w image was corrected for intensity non-uniformity with
281 N4BiasFieldCorrection (Tustison et al., 2010) distributed with ANTs 2.3.3 (Avants et al.,
282 2008) and used as T1w-reference throughout the workflow. The T1w-reference was
283 then skull-stripped using ANTs workflow with OASIS30ANTs as target template. Brain
284 tissue segmentation of gray-matter, white-matter, and cerebrospinal fluid was performed
285 on the brain-extracted T1w using fast (FSL 6.0.5.1:57b01774, RRID:SCR_002823),
286 (Zhang et al., 2001). Brain surfaces were reconstructed using recon-all from FreeSurfer
287 version 7.2.0 (Dale et al., 1999)), and the brain mask estimated previously was refined
288 with a custom variation of the method to reconcile ANTs-derived and FreeSurfer-derived
289 segmentations of subcortical gray matter including striatal subregions (Fischl et al.,
290 2002). Volume-based spatial normalization to one standard space was performed
291 through nonlinear registration, using brain-extracted versions of both the T1w reference
292 and T1w template. *FSL's MNI ICBM 152 non-linear 6th Generation Asymmetric*
293 *Average Brain Stereotaxic Registration Model* (Evans et al., 2012), RRID:SCR_002823;
294 TemplateFlow ID: MNI152NLin6Asym] was selected for spatial normalization.

295

296 *Functional preprocessing*

297 A skull-stripped reference volume was generated using a custom methodology of
298 *fMRIPrep*. Head-motion parameters were estimated using mcflirt (FSL
299 6.0.5.1:57b01774, (Jenkinson et al., 2002). The estimated fieldmap was aligned with
300 rigid-body registration to the target EPI reference run. Field coefficients were mapped
301 onto the reference EPI using the rigid-body transform. BOLD runs were slice-time
302 corrected using 3dTshift from *AFNI* (Cox & Hyde, 1997) RRID:SCR_005927). The
303 BOLD reference images were co-registered to the T1w reference using bbrregister
304 (*FreeSurfer*, (Greve & Fischl, 2009), with six degrees of freedom. Noise regressors
305 were estimated based on the preprocessed BOLD. Framewise displacement (FD) was
306 computed using two formulations following Power (Power et al., 2014), and Jenkinson
307 (Jenkinson et al., 2002). Physiological regressors were extracted from eroded
308 cerebrospinal fluid and white-matter volumes for use in subsequent, component-based
309 noise corrections (*CompCor*, Behzadi et al. 2007). The BOLD time-series were
310 resampled into standard space in a single interpolation step by composing all the
311 pertinent transformations (i.e., head-motion transform matrices, susceptibility distortion
312 correction, and co-registrations to anatomical and output spaces). Volumetric
313 resamplings were performed using ANTs, configured with Lanczos interpolation to
314 minimize the smoothing effects of other kernels (Lanczos 1964).

315

316 **Statistical analyses**

317 Statistical analyses were conducted in R (version 4.2.2). Raw data, code, and
318 extended analyses and supplementary tables are available at the following link:

319 <https://osf.io/pqtby/>. Unless stated otherwise, linear mixed effects regressions or general
320 linear mixed effects regressions were employed for data that repeated within
321 participants (e.g., single-trial RTs). Mixed effects models included random intercepts for
322 each participant and random slopes for fixed effects that repeated within participants. In
323 case of non-converging models, we followed recommendations in (Brown, 2021),
324 iterating through the following until they converged: (1) using the "bobyqa" optimizer, (2)
325 increasing the number of iterations, (3) forcing zero correlations among random effects,
326 and (4) dropping random effects based on model comparison. RTs that fell 3 absolute
327 deviations from an individual's median RT were excluded ($n=2$ on average per
328 participant). Measures were mean centered within or across participants or effect-coded
329 prior to model fitting.

330

331 *Calculating trial-by-trial shifts in the moving average reward rate*

332 We computed an exponentially weighted moving average (EWMA) of rewards
333 and losses across trials (\$1, \$0.5, or \$0; See *Figure 1C-D* for a schematic). Each trial
334 was assigned a value based on the recent reward and loss history. High values
335 indicated more gains than losses, whereas low values indicated more losses than gains.
336 We used an exponentially weighted (rather than simple) moving average to emphasize
337 recent time points, which have a larger impact on psychological state, while still
338 incorporating data points from farther in the past (Awheda & Schwartz, 2016). We used
339 the following update rule:

340

341
$$EWMA_t = \alpha \times r_t + (1 - \alpha) \times EWMA_{t-1}$$

342

343 In this formula, the $EWMA_t$ represents the EWMA at the current trial, t , α is the
344 smoothing factor or learning rate parameter that determines the influence of the most
345 recent observation on the moving average, and r represents the reward on the current
346 trial, t . To prevent disproportionate initial weighting and to ensure the average reward
347 rate stabilized, the first 3 trials were omitted from the EWMA measure. To balance
348 recent and historical data, α was based on an 8-period span, calculated as: $\alpha = 2/(N+1)$,
349 where N was set to 8 to match the number of trials in each fixed reward and loss block.
350 However, we found that using EWMA's derived from smoothing factors of 5- and 10-
351 period spans did not alter the pattern of results.

352 As an exploratory analysis, we also tested whether individual differences in
353 optimal learning rates for the average reward rate variable differed by SES. To do so,
354 we fit a model that estimated the learning rate as a free parameter for each participant
355 using R's base *optim* function with the L-BFGS-B algorithm. The algorithm identified the
356 learning rate per participant that minimized that residual sum of squares (RSS) in a
357 model predicting subsequent RTs from the Exponentially Weighted Moving Average of
358 Reward for each participant.

359 *Characterizing behavioral responses to rewards*

360 We examined how immediate feedback (win versus loss outcomes) and
361 fluctuations in the average reward rate shaped RTs and guesses. We first fit a model

362 predicting RTs from the preceding trial's feedback (win, loss), the moving average of
363 reward, and their interaction. We then re-fit this model after adding SES as a covariate
364 and interaction term. We also examined the influence of immediate reward and average
365 reward rate fluctuations on choices—specifically, how likely an individual was to repeat
366 their prior guess or switch to a different guess (i.e., switched or stayed). Therefore, the
367 dependent variable was whether an individual had repeated their prior choice (switched
368 = 1; stayed = 0) and the independent variables were the preceding trial feedback (win,
369 loss), the moving average of reward, and their interaction. We re-fit this model after
370 adding SES as a covariate and interaction term. All models included trial number as a
371 covariate to control for the general effects on time on task. Since there were only 8
372 neutral trials per participant across the task, trials that followed neutral feedback were
373 excluded from analysis.

374

375 **The relationship between SES and striatal volumes**

376 Three linear mixed effects models were fit to examine the association between
377 SES and ROI volumes, separately for the caudate, putamen, and nucleus accumbens.
378 Each model predicted volume from SES, hemisphere, and their interaction, to determine
379 whether the influence of SES was stronger for one particular hemisphere. Age, sex, and
380 intracranial volume were also included as covariates. ROIs with volumes that fell > 3
381 absolute deviations from the sample median were excluded (all regions for 1 participant,
382 the caudate and right nucleus accumbens for another).

383

384 **Examining reward-driven striatal responses to reward and average reward rate**
385 **fluctuations across time**

386 To ascertain if striatal activations during reward differed from loss blocks and to
387 examine their covariance with average reward rate fluctuations, we conducted
388 neuroimaging analyses with Nilearn. The scripts and data are publicly accessible
389 (<https://osf.io/pqtbby/>). The approach involved two separate general linear models
390 (GLMs) applied to participant data within MNI coordinate space. The first model had
391 distinct regressors for reward and loss conditions. The second model incorporated a
392 regressor for the exponentially weighted moving average of reward, resampled at the
393 fMRI's TR. Both models were convolved with SPM's hemodynamic response function
394 and controlled for head movement and noise components (3 translation and rotation
395 parameters, plus top 5 principal AcompCor components (defined in a combined white
396 matter and cerebrospinal fluid mask). This analysis yielded z-value effect size maps for
397 each subject. The maps were entered into a group-level analysis to identify striatal
398 voxels that were sensitive to the distinct reward versus loss blocks or to the average
399 reward rate. Sensitivity was defined by voxel significance within the anatomical striatal
400 mask from the Harvard-Oxford Atlas (FDR-corrected $p < 0.05$, minimum cluster size of
401 10). For each analysis, we calculated the mean z value per participant across
402 responsive voxels, separately for the caudate, putamen, and nucleus accumbens in
403 each hemisphere. Participants therefore had 6 z values (1 per ROI) for each analysis.
404 These values represented the average effect size for the differences in activations
405 between reward and loss blocks and the relationship with the average reward rate.

406 To assess the degree to which these effect sizes deviated from zero, we fit 2
407 intercept-only linear mixed-effects models, predicting mean z values per ROI, controlling
408 for age and sex, with random intercepts per participant to account for repeated
409 measures across hemispheres. We excluded outlier values that fell 3 absolute
410 deviations from the sample's median (1 value for the left putamen and 1 for the left
411 caudate). Including outliers did not change the pattern of results.

412 **Examining how reward-driven striatal responses differ by SES**

413 Finally, we tested how SES related to activation level differences between reward
414 and loss blocks, as well as the degree to which striatal activations covaried with
415 fluctuations in the average reward rate. To this end, we fit 2 linear mixed effects models.
416 The dependent variables were z values reflecting, either, activation level differences for
417 reward and loss blocks or the covariance between striatal activations and average
418 reward rate fluctuations. Both models included SES, hemisphere, and their interaction
419 as independent variables, and covariates for age and sex.

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Results

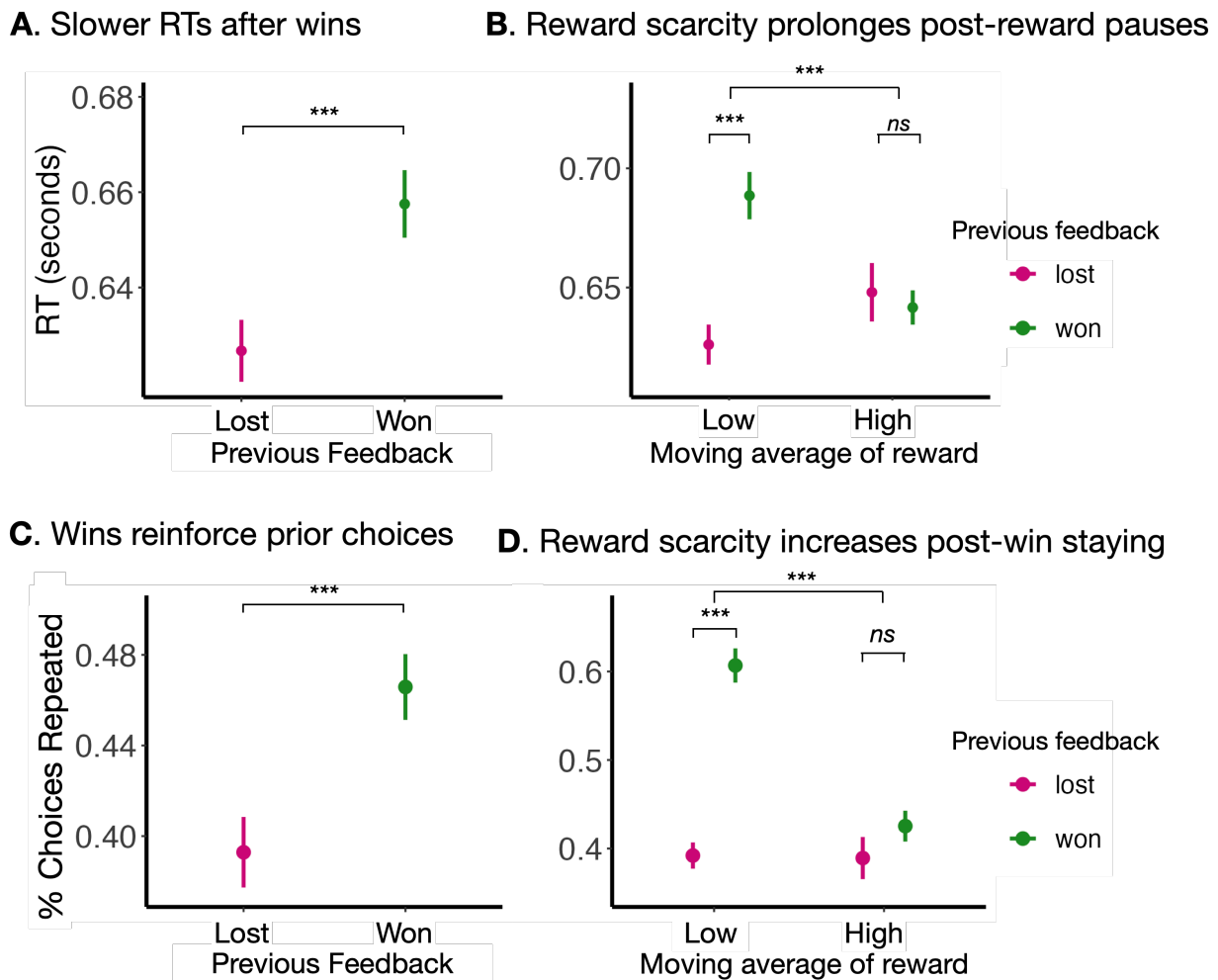
429 We first describe how behavioral responses, specifically response times (RT)
430 and choices, are influenced by immediately rewarding outcomes, and covary with
431 fluctuations in the average reward rate across time. We then describe how these
432 behavioral responses differ by SES. Turning to the neuroimaging data, we then explore
433 the association between SES and the volume of the putamen, caudate, and nucleus
434 accumbens. Furthermore, we examine differences in striatal activations during reward
435 versus loss blocks, and examine how these activations covary with temporal
436 fluctuations in the average reward rate. Finally, we focus on disparities in striatal
437 responses across SES.

438

439 **Average reward rate fluctuations influence RTs and post-reward pausing**

440 Adolescents responded more slowly after winning than losing (i.e., post-reward
441 pausing: $\beta = 0.02$, $SE = 0.005$, $t(262) = 4.68$, $p < .001$; Figure 2A). Furthermore, trial-by-
442 trial RTs covaried with fluctuations in the average reward rate, such that a higher
443 average reward rate led to faster RTs ($\beta = -0.04$, $SE = 0.02$, $t(98) = -2.32$, $p = .022$).
444 Fluctuations in the average reward rate also interacted with immediate feedback to
445 shape RTs: periods of reward scarcity amplified post-reward pausing (*Reward Rate x*
446 *Preceding feedback*: $\beta = -0.05$, $SE = 0.02$, $t(104) = -3.53$, $p < .001$, Figure 2B),
447 indicating responses to immediate reward were amplified by a history of low rewards. In
448 fact, post-reward pausing was only observed when rewards were scarce but not when
449 they were plentiful (*effect of preceding feedback when the reward rate is centered at -*

450 1SD below the mean: $\beta = 0.04$, $SE = 0.007$, $t(100) = 6.23$, $p < .001$; above the mean: β
 451 $= 0.003$, $SE = 0.008$, $t(104) = 0.43$, $p = .665$). These findings show that adolescents
 452 tracked fluctuations in the average reward rate, which shaped response times across
 453 time and modulated sensitivity and responses to immediate reward.
 454



455

456 **Figure 2. Immediate reward and the moving average reward rate shape RTs and**
 457 **choices.** (A) Adolescents responded more slowly after a win than a loss ($p < 0.001$). (B)
 458 A lower average reward rate amplified post-reward slowing (*interaction*: $p < 0.001$)
 459 indicating heightened behavioral responses to reward when rewards were scarce. (C)
 460 Adolescents were most likely to repeat a guess when their guess had been rewarded on
 461 the previous trial ($p < 0.001$). (D) These “post-reward stay” effects were amplified by a
 462 low average reward rate ($p < 0.001$). In all figures, the mean and within-subject error
 463 bars are plotted. The moving average reward rate was divided into low and high

464 average reward rate using a median split for ease of visualization. Note that we model
465 the average reward rate continuously in all analyses.
466

467

468 **Average reward rate fluctuations influence guess switching**

469 Immediate feedback reinforced decisions on subsequent trials: when adolescents
470 won, they were more likely to repeat their prior guess on the subsequent trial than if
471 they had lost ($\beta = -0.31$, $SE = 0.04$, $z = -7.31$, $p < .001$; *Figure 2C*). A lower average
472 reward rate also increased the likelihood of repeating a previously rewarded guess (i.e.,
473 increased win-stay, lose-switch effects (Reward Rate x Preceding Feedback: $\beta = 0.57$,
474 $SE = 0.12$, $z = 4.73$, $p < .001$; *Figure 2D*). Indeed, win-stay effects were most prominent
475 when the average reward rate was low, indicating a history of low rewards increased the
476 tendency to stick with a rare rewarding option (*main effect of immediate feedback on*
477 *choices when the average reward rate is centered at -1SD below the mean: $\beta = -0.51$,*
478 *$SE = 0.05$, $z = -9.64$, $p < .001$; above the mean: $\beta = -0.11$, $SE = 0.06$, $z = -1.75$, $p =$*
479 *.080*). In general, a history of high rewards (a higher average reward rate) also increased
480 the likelihood of switching guesses across trials ($\beta = 0.68$, $SE = 0.15$, $z = 4.46$, $p <$
481 $.001$), suggesting a greater tendency to make alternative exploratory decisions when
482 rewards were abundant. These findings suggest that a history of low reward increases
483 the tendency to stick with a previously rewarding option, and reduces the tendency to
484 explore alternatives for reward.

485

486 **Reward rate fluctuations influence post-reward pausing more in higher-SES**
487 **adolescents**

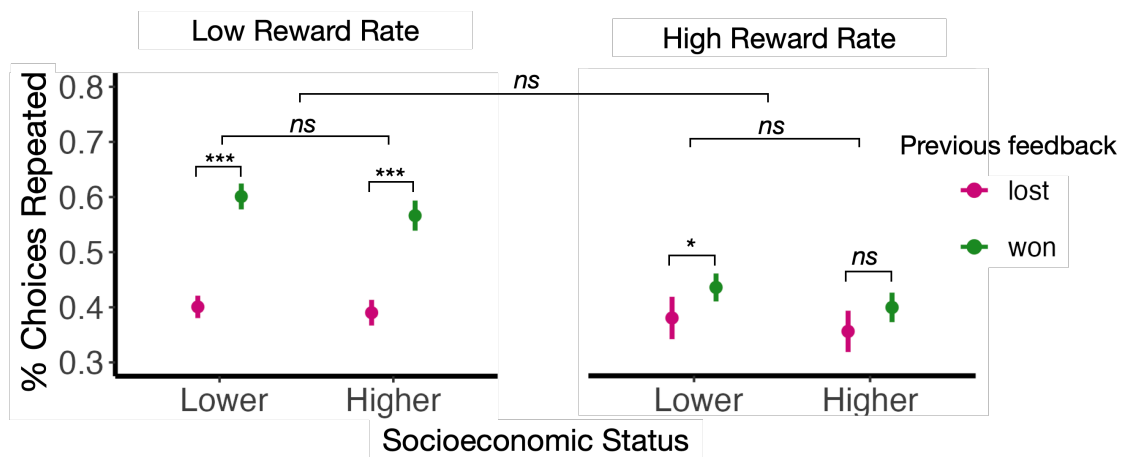
488 Immediate reward and average reward rate fluctuations influenced choices
489 similarly regardless of SES (*SES x feedback*: $\beta = 0.09$, $SE = 0.05$, $z = 1.73$, $p = .084$;
490 *SES x average reward rate*: $\beta = -0.25$, $SE = 0.18$, $z = -1.40$, $p = .161$; *SES x feedback*
491 *type x moving average*: $\beta = 0.18$, $SE = 0.15$, $z = 1.21$, $p = .225$; *Figure 3A*). Additionally,
492 these distinct temporal dimensions of reward influenced RTs similarly, regardless of
493 SES (*SES x feedback*: $\beta = -0.002$, $SE = 0.005$, $t(109) = -0.45$, $p = .651$; *SES x average*
494 *reward rate*: $\beta = 0.02$, $SE = 0.02$, $t(100) = 1.04$, $p = .301$; *Figure 3A*).

495 However, reward-rate fluctuations modulated post-reward pausing more in
496 higher- than lower-SES adolescents (*SES x feedback type x moving average*: $\beta =$
497 -0.04 , $SE = 0.02$, $t(105) = -2.54$, $p = .013$; *Figure 3B*). That is, higher-SES adolescents
498 slowed more following rare rewards (main effect of SES when the reward rate is
499 *centered at -1SD below the mean to reflect reward scarcity*: $\beta = 0.02$, $SE = 0.007$, $t(684)$
500 $= 2.20$, $p = .028$; *Figure 3B*). When rewards were plentiful, higher-SES adolescents
501 slowed less following rewards than lower-SES adolescents (*centered at +1SD above*
502 *the mean to reflect reward abundance*: $\beta = -0.02$, $SE = 0.008$, $t(4949) = -2.05$, $p =$
503 $.041$; *Figure 3B*) though neither group showed significant evidence of post-reward
504 pausing when rewards were plentiful ($ps > 0.087$). Interestingly, SES was unrelated to
505 individual differences in optimal learning rates ($\beta = 0.06$, $SE = 0.04$, $t(114) = 1.32$, $p =$
506 $.189$), suggesting that heightened post-reward pausing was not driven by a greater
507 tendency to update expectations in response to new information. These findings

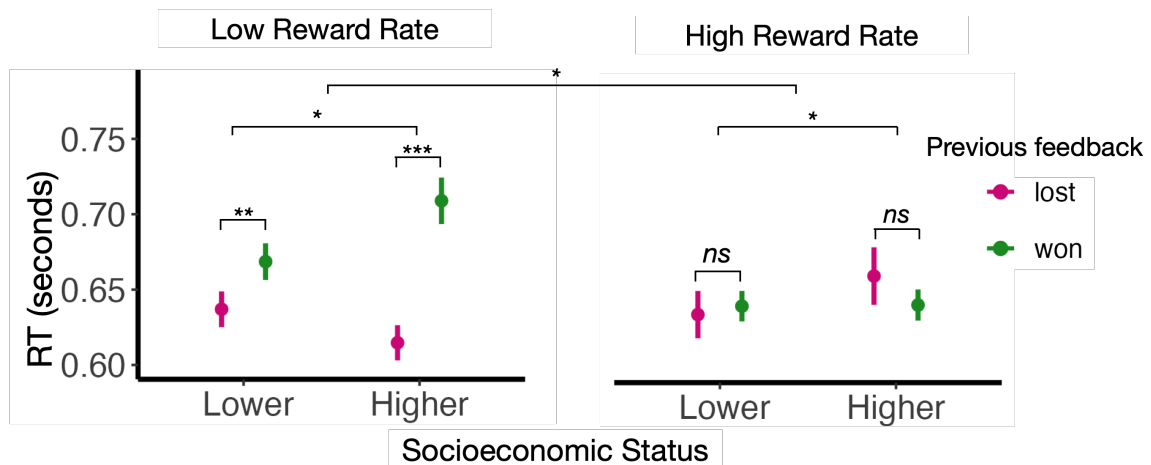
508 suggest that adolescents from lower-SES backgrounds were less likely to adapt
 509 responses to immediate reward based on average reward rate fluctuations. Analyses
 510 reported in our extended analyses on the Open Science Framework
 511 (<https://osf.io/9vhtw>) demonstrate these results are robust when using education and
 512 income to separately characterize SES.

513

A. The influence of reward on choices does not differ by SES



B. Reward rate fluctuations shape post-reinforcement pausing differently by SES



514

515 **Figure 3. Reward rate fluctuations modulate post-reward pausing more in higher-**
 516 **SES adolescents.** (A) Reward rate fluctuations influenced choice switching following
 517 immediate reward similarly by SES. (B) Reward rate fluctuations modulated post-reward

518 pausing more among higher- than lower-SES adolescents. This led to greater RT
519 slowing following rewards when rewards were scarce among higher-SES adolescents.
520 In all figures, we depict the mean and within-subject error bars. The moving average
521 reward rate and socioeconomic status were divided into low and high bins using a
522 median split for ease of visualization. Note that we model these variables continuously
523 in the analyses reported in the paper.
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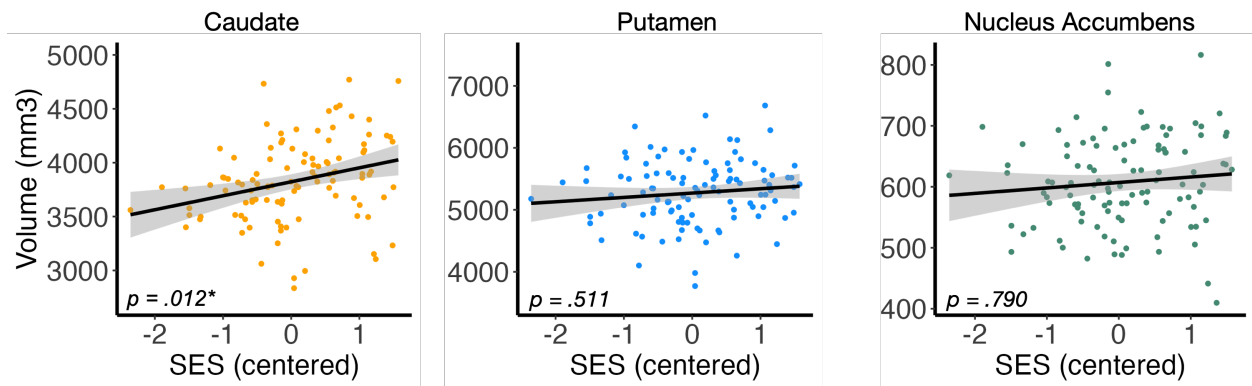
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526 Lower SES correlates with smaller caudate volumes

527 Higher SES was associated with larger caudate volumes ($\beta = 96.61$, $SE = 37.65$,
528 $t(103) = 2.57$, $p = .012$; *Figure 4*). In contrast, there were no significant associations
529 between SES and the volumes of the putamen ($\beta = 32.78$, $SE = 49.74$, $t(103) = 0.66$, p
530 $= .511$) or nucleus accumbens ($\beta = 1.94$, $SE = 7.26$, $t(104) = 0.27$, $p = .790$). Moreover,
531 there were no SES x hemisphere interactions in any ROI (*all ps* > 0.590), demonstrating
532 that SES-related differences in volumes did not differ by hemisphere.

533

Lower SES correlates with smaller caudate volumes



534

535 **Figure 4. SES and striatal volumes.** SES positively correlated with the volume of the
536 caudate ($p = 0.012$), but not the putamen or nucleus accumbens ($ps > 0.510$). For ease
537 of visualization, data points reflect the average volume of the left and right hemisphere,
538 but the statistics reported in the text are from models that treat the left and right

539 hemisphere as repeated measures within participants. Individual data points represent
540 participant level data, and the grey shading reflects the standard error of the mean.
541

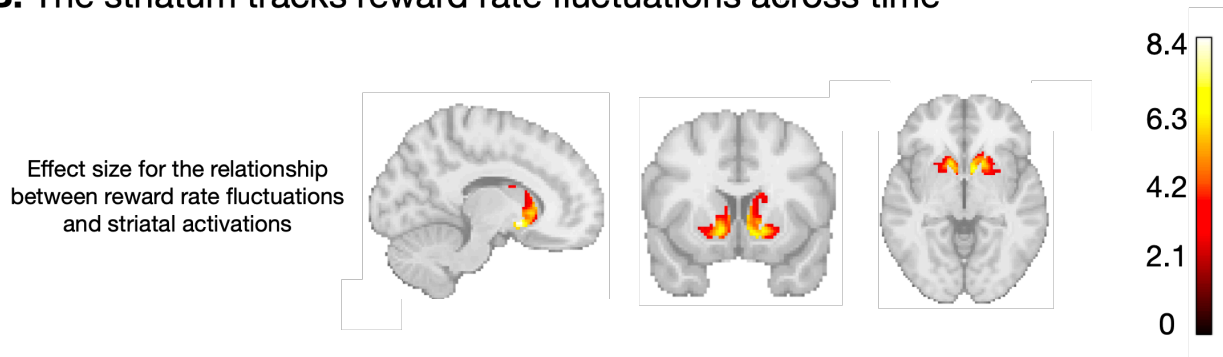
542 **The striatum tracks fluctuations in the average reward rate**

543 Across adolescents, mean activations were larger during reward than loss blocks in the
544 caudate ($\beta = 0.50$, $SE = 0.07$, $t(106) = 6.74$, $p < .001$), putamen ($\beta = 0.61$, $SE = 0.08$,
545 $t(106) = 7.83$, $p < .001$), and nucleus accumbens ($\beta = 0.77$, $SE = 0.08$, $t(107) = 9.87$, p
546 $< .001$; *Figure 5A*). Furthermore, striatal activations covaried with the average reward
547 rate, such that a higher average reward rate led to greater activations in the caudate (β
548 $= 0.77$, $SE = 0.08$, $t(105) = 9.60$, $p < .001$), putamen ($\beta = 0.66$, $SE = 0.07$, $t(105) = 8.94$,
549 $p < .001$), and nucleus accumbens ($\beta = 1.32$, $SE = 0.09$, $t(103) = 14.07$, $p < .001$; *Figure*
550 *5B*). These findings show that the striatum not only responds more to reward than loss in
551 general, but tracks moment-by-moment shifts in the average reward rate across time.
552

A. Higher striatal activations in reward than loss blocks



B. The striatum tracks reward rate fluctuations across time



553

554 **Figure 5. Striatal activations covary with reward and loss blocks and reward rate**
555 **fluctuations across time.** (A) Across the sample, activations in the striatum were
556 greater during reward than loss blocks ($ps < 0.001$). (B) Moreover, the striatum tracked
557 average reward rate fluctuations across time, even after FDR correction for multiple
558 comparisons across every voxel in the striatum ($ps < 0.001$). In A and B, the color bar
559 represents z-values and color intensity reflect the strength of the effect.

560

561

562 Lower SES correlates with reduced striatal responses to reward

563 Lower SES correlated with smaller activation level differences between reward

564 and loss blocks in the caudate ($\beta = 0.22$, $SE = 0.09$, $t(105) = 2.54$, $p = .013$) and

565 putamen ($\beta = 0.25$, $SE = 0.09$, $t(104) = 2.73$, $p = .007$) and marginally in the nucleus

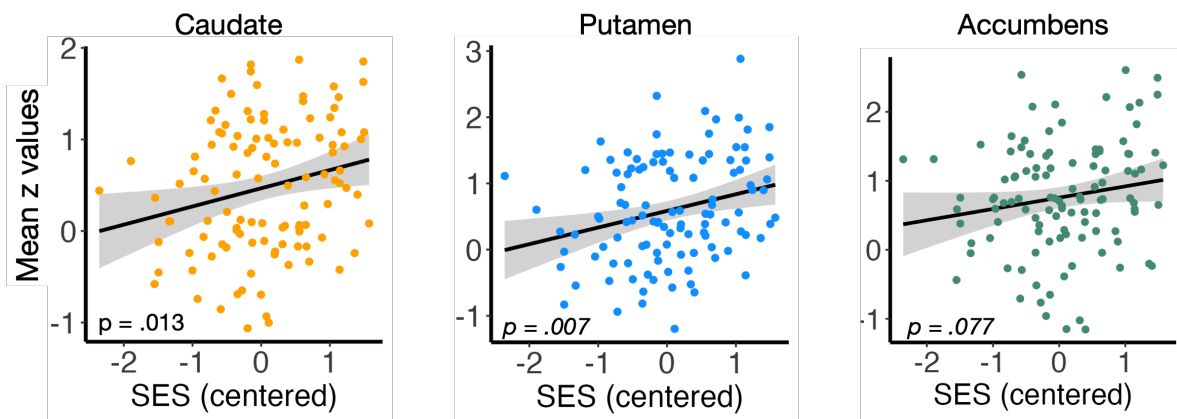
566 accumbens (marginal effect: $\beta = 0.16$, $SE = 0.09$, $t(106) = 1.79$, $p = .077$; Figure 6A).

567 None of these effects differed by hemisphere ($SES \times hemisphere$: all $ps > 0.29$).

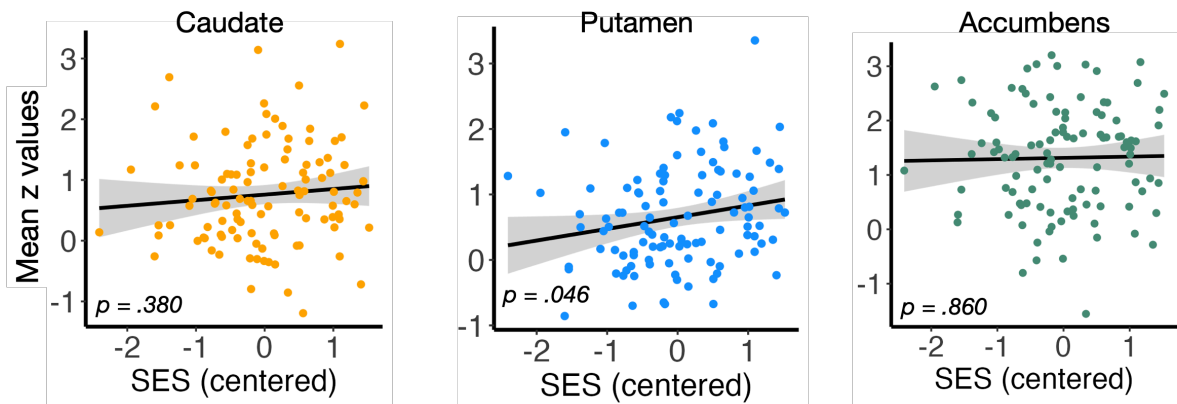
568 Furthermore, striatal activations covaried with average reward rate fluctuations more

569 strongly in higher-SES adolescents in the putamen ($\beta = 0.17$, $SE = 0.09$, $t(104) = 2.02$,
 570 $p = .046$; Figure 6B), but not the caudate ($\beta = 0.08$, $SE = 0.10$, $t(104) = 0.88$, $p = .380$)
 571 or nucleus accumbens ($\beta = 0.02$, $SE = 0.11$, $t(101) = 0.18$, $p = .860$). None of these
 572 effects differed by hemisphere ($SES \times hemisphere$: all $ps > 0.21$). Of note, the
 573 relationship between SES and reward-driven activations also did not differ by striatal
 574 subregion ($SES \times subregion$ interaction: all $ps > 0.10$).
 575

A. Lower SES correlates with smaller reward-driven activations



B. Putamen activations track reward rate fluctuations less in lower-SES adolescents



576

577 **Figure 6. Lower SES correlates with reduced striatal activations to reward (A)**
 578 Lower SES correlated with less activation differences between reward and loss blocks
 579 in the caudate and putamen and marginally in the nucleus accumbens. (B) The

580 relationship between striatal activations and average reward rate fluctuations was
581 stronger for higher- than lower-SES adolescents in the putamen ($p=0.046$), but not in
582 the caudate or nucleus accumbens ($p>0.37$). For ease of visualization, individual data
583 points reflect mean z values across the left and right hemisphere. Statistics reported in
584 the text model the left and right hemisphere separately as repeated measures.
585 Individual data points represent participant data, and the grey shading reflects the
586 standard error of the mean.

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Discussion

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We asked how socioeconomic status in adolescence was related to reward-driven responses in the brain and behavior. Drawing on influential models of decision-making (Constantino & Daw, 2015; Niv et al., 2006, 2007), we examined how choices, response times, and striatal activations were shaped by immediate reward outcomes and previous reward history (average reward rate fluctuations across time). We found that, behaviorally, participants were more likely to repeat a guess if it had led to a win (win-stay, lose-switch effects) and responded more slowly after receiving a reward (post-reward pausing). Fluctuations in the average reward rate also shaped behavior: a higher reward rate hastened RTs and increased guess switching. Moreover, a low reward rate increased behavioral sensitivity to immediately rewarding outcomes, augmenting win-stay, lose-switch effects and post-reward pausing. Notably, compared to higher-SES adolescents, lower-SES adolescents exhibited reduced post-reward pausing when rewards were scarce. We also observed that across participants, striatal activations were larger during reward than loss blocks, and covaried with fluctuations in the average reward rate across time. However, relative to higher-SES adolescents, lower-SES adolescents displayed reduced activations during reward relative to loss

606 blocks in the caudate and putamen, and marginally in the nucleus accumbens. And,
607 putamen activations tracked average reward rate fluctuations less in lower-SES
608 adolescents. These findings show that the striatum tracks average reward rate
609 fluctuations, which shape choices and response times (Hamid et al., 2016; Niv et al.,
610 2006, 2007; Wang et al., 2013, 2021). They also link lower SES in adolescence to
611 reduced reward sensitivity, both in the brain and behavior.

612 We found that adolescents tracked fluctuations in the average reward rate across
613 time, which influenced decisions and response times. When rewards were abundant,
614 individuals were more likely to switch choices across trials. These findings align with
615 studies in human adults (Constantino & Daw, 2015; Niv et al., 2007; Sukumar et al.,
616 2021) and support theories of decision making (Constantino & Daw, 2015; Sukumar et
617 al., 2021). These theories argue that when the average environmental reward rate is
618 *lower* than an option's perceived value, it is rational to 'stay' with a rewarding option due
619 to the limited prospects of finding rewards elsewhere. Conversely, when the
620 environmental reward rate is *higher* than the perceived value of an option, it makes
621 sense to switch to explore alternative sources of reward. It is possible, then, that
622 adolescents used the average reward rate as a threshold for whether to switch or stay
623 with a previous choice. Future research could examine how the tendency to track
624 average reward rate fluctuations develops—and whether adolescents—given their
625 heightened sensitivity to reward (Cohen et al., 2010; Davidow et al., 2016; Galvan,
626 2010; Galvan et al., 2006) might be even more attuned to fluctuations in the average
627 reward rate across time than adults.

628 A higher average reward rate also covaried with faster RTs. This finding is
629 consistent with research in human adults (Beierholm et al., 2013; Otto & Daw, 2019)
630 and supports theories arguing that fluctuations in the average reward rate shape the
631 cost time (Niv et al., 2006, 2007). That is, when rewards are abundant, action delays
632 are presumably more costly because one forfeits relatively more potential rewards,
633 incentivizing faster responses. Interestingly, other researchers have theorized that
634 rewards also govern the opportunity cost engaging effort and sustaining attention
635 (Esterman et al., 2016; Esterman & Rothlein, 2019; Kurzban et al., 2013; Lin et al.,
636 2022; Massar et al., 2016; Otto & Daw, 2019) raising the possibility that average reward
637 rate fluctuations shape diverse aspects of cognition—such as fluctuations in attention
638 (Decker et al., 2023, 2022; Decker & Duncan, 2020). Our findings therefore not only
639 support theories linking reward rate fluctuations to motivation and decision-making and
640 extend these ideas to human adolescents, but raise questions about the influence of
641 reward rate fluctuations on other aspects of cognition.

642 Adolescents were also responsive to immediately rewarding outcomes, in line
643 with previous research (Hamid et al., 2016; Reynolds et al., 2001): they were most likely
644 to repeat a previous choice if it had led to a reward on the prior trial and responded
645 more slowly after a reward outcome, a phenomenon known as “post-reward pausing”
646 (Crossman, 1968; Felton & Lyon, 1966; McMillan, 1971; Schlinger et al., 2008; Wallace
647 & Mulder, 1973; Williams et al., 2011). Notably, these effects were amplified by a lower
648 average reward rate. Our finding adds to a growing body of research suggesting the
649 background average reward rate modulates sensitivity to immediate reward. Indeed, in

650 animals and humans, post-reward pausing is prolonged when rewards are scarce
651 (Schlinger et al., 2008). Furthermore, fewer recent rewards and lower tonic dopamine
652 amplify phasic dopamine firing (Hamid et al., 2016)—a finding that potentially provides a
653 neurobiological explanation for the increased rewards responsivity we observed here
654 when reward were scarce. Slower responses after unexpected reward could also reflect
655 surprise due to the infrequency of the event (Decker, Finn, et al., 2020) or heightened
656 response caution that facilitated more deliberate decision-making (Schlinger et al.,
657 2008, p. 50). Altogether, this finding shows that average reward rate fluctuations
658 influenced responses to immediate outcomes, which shaped choices and RTs. When
659 adolescents tune into the average environmental reward rate, they may make more
660 adaptive decisions according to the overall rewards available in the environment.

661 We also observed that the extent of RT slowing after rare rewards varied by
662 SES. Adolescents from higher-SES backgrounds showed greater post-reward pausing
663 than lower-SES adolescents when rewards were scarce. This finding could reflect
664 greater attunement to reward rate fluctuations among higher-SES adolescents, which
665 would be expected to increase the saliency of receiving a rare reward when the reward
666 rate was low. However, exploratory analyses showed that SES did not correlate with
667 learning rates—the tendency to update the average reward rate in response to new
668 outcomes. Thus, greater post-reward pausing may instead reflect a greater responsivity
669 to rewards in reward-scarce contexts specifically, rather than a general tendency to
670 more readily update the average reward rate.

671 Interestingly, reward rate fluctuations covaried with striatal activations in the
672 caudate, putamen and nucleus accumbens, such that a higher reward rate led to
673 greater activations in these regions. These findings are consistent with animals studies
674 showing that tonic dopamine fluctuations in the striatum track the average reward rate
675 and in doing so shapes motivational vigor and decision-making (Hamid et al., 2016;
676 Wang et al., 2013, 2021), and, as far as we know, is the first human fMRI study
677 demonstrating this relationship.

678 Our results extend prior findings linking lower SES to diminished reward
679 sensitivity in *neocortical* regions like the anterior cingulate cortex (Palacios-Barrios et
680 al., 2021) and parietal cortex (White et al., 2022). Indeed, we observed that the extent of
681 reward-driven activations in the striatum differed by SES. Higher-SES adolescents
682 showed greater reward-driven activations than lower-SES adolescents in the putamen,
683 caudate, and marginally in the nucleus accumbens. Moreover, putamen activations
684 tracked average reward rate fluctuations less in lower-SES adolescents. Notably, prior
685 studies employed incremental learning tasks in which adolescents learned the value of
686 cues in predicting outcomes over time. Our focus on a reward task that did *not* involve
687 learnable cue-outcome contingencies broadens the literature by showing that reduced
688 reward sensitivity is even observed when eliminating learning demands.

689 Our findings support proposals that lower SES environments reduce reward
690 sensitivity (Seligman, 1972). Past literature suggests that chronic stress diminishes the
691 belief that actions have consequences rendering individuals less motivated to pursue
692 rewarding outcomes (Seligman, 1972). It is therefore possible that chronic stress and

693 reduced perceived control, which are more common among lower-SES individuals
694 (Farah, 2018; Hackman et al., 2010; Hackman & Farah, 2009; McLaughlin et al., 2014)
695 mediated the effects we observed here. Targeted research that employs direct measures
696 of stress could directly test this mechanism.

697 The present findings offer insights into why cognitive performance (Noble et al.,
698 2007) and emotional well-being (Reiss, 2013) are often reduced in lower SES
699 adolescents. Reward sensitivity plays a vital role in many aspects of cognition,
700 influencing everything from the ability to learn important associations (Davidow et al.,
701 2016) to the ability to remain attentive to important events (Esterman & Rothlein, 2019;
702 Shenhav et al., 2013). Rewards boost motivation (Frömer et al., 2021; Schultz, 1993;
703 Westbrook et al., 2021; Westbrook & Braver, 2016), and support success in short and
704 long-term endeavors, such as academic and workplace pursuits. Disparities in reward
705 sensitivity, therefore, may contribute to disparities in learning, attentional performance,
706 and motivation. Given the intimate link between reward sensitivity and emotional well-
707 being, reduced reward sensitivity may contribute to the higher rates of depression
708 (Auerbach et al., 2022; Reiss, 2013) and lower life satisfaction observed in lower-SES
709 groups (Kahneman & Deaton, 2010). On a broader level, these insights stress the
710 importance of socioeconomic policies (Farah, 2018) aimed at reducing the burdens of
711 poverty to foster cognitive and emotional well-being in society.

712

713 **Data availability statement:** Code and data can be found at the following link:

714 <https://osf.io/pqtby/>

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Methodology: A.D., C.C.C.B.; Project Administration: N.A.H., A.I., M.A.G., H.G., Y.C.T.,
R.R.; Resources: J.D.E.G.; Software: A.D., S.L.M, C.C.C.B.; Supervision: J.D.E.G.;
Visualization: A.D.; Writing – Original Draft: A.D.; Writing – Review and Editing: all
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